

Managing Right-Wing Challengers Through International Migration Cooperation

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Abstract

The electoral rise of right-wing populism has reshaped domestic competition across Western democracies. I develop a theory of *foreign policy co-optation* that explains when and why governments appropriate flexible international cooperation instruments central to the narrative of the opposition to reduce their electoral threat. As right-wing populist parties gain seats, mainstream governments co-opt bilateral migration instruments to counter domestic political challengers. These effects are especially pronounced for governments most challenged by right-wing narratives, rather than ideologically adjacent governments. I employ an original dataset of over 300 bilateral *cooperation arrangements on migration* (CAMs) signed by ten Western frontier states with partners across five continents over the last thirty years. The findings suggest that right-wing populist pressure enables executives to manage electoral opposition through foreign policy, highlighting the need to revisit assumptions about the domestic sources of international cooperation and migration policy. (142)

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Introduction

The electoral rise of right-wing populist and radical-right parties has reshaped political competition across Western industrialized states over the last thirty years. At the same time, democratic governments have spearheaded a number of security-oriented bilateral migration agreements emphasizing border management and migration control. On January 31, 2020, for instance, the Interior Minister of Italy and the Minister of Security and Civil Protection of Côte d’Ivoire finalized an accord to further expand joint-police training and technical assistance to combat irregular migration (UniTO 2020).¹ Two years later, Spain’s Interior Minister renewed security cooperation with Mauritania, highlighting joint efforts to curb departures to the Canary Islands (La Moncloa 2022). More recently, on August 20, 2025, the United Kingdom’s Security Minister and Iraq’s Deputy Foreign Minister signed a border cooperation and irregular migration management agreement, committing nearly 1 million pounds to joint law enforcement training on border security (UK Government, Home Office & FCDO 2025).² Such collaboration is often attributed to multilateral strategies or migrant pressures at the border given that bilateral agreements can be effective tools to manage migration as an internal and external security concern (Rudolph 2003).

Studies have indeed noted that Western democracies attempt to control migrants abroad to evade the domestic political and practical burden of caring for them at home,³ with an expectation that right-leaning governments would restrict more often. Yet, the propensity of such bilateral cooperation agreements varies across migrant pressure and multilateral initiatives, while governments implementing these policies also vary across the political spectrum.

1. Recently, in August 3, 2026, Italy and Algeria agreed to deepen cooperation on irregular migration and cross-border crime as well (Italian Government 2025; Anadolu Agency 2026).

2. Government spending in the area of border security and migration is expected to grow as much as 81 billion USD by 2030. In the EU alone, the border guard agency Frontex recently took the highest budget of all EU agencies with 5.6 billion Euros for their 2021-27 budget cycle (Akkerman 2023).

3. Refer to Hollifield (1992), Huysmans (2000), Lahav (2004), Lavenex (2006), Godenau et al. (2008), Andersson (2014), Gammeltoft-Hansen (2013), Menjívar (2014), Wong (2015), Mann (2016), FitzGerald (2019), Kent, Norman, and Tennis (2020), FitzGerald (2020), Money (2021), Money and Lockhart (2019), Léonard and Kaunert (2020), Bello (2022), Panebianco (2022), Pacciardi and Berndtsson (2022), Fakhoury (2022), Lavenex and Piper (2022), Brettell and Hollifield (2022), Cassarino, Gabrielli, and Perrin (2023), Norman and Micinski (2023), and Lahav and Messina (2024).

If we then focus on the agreements themselves, despite extensive work on migration management, the determinants of such publicized and direct security tool are split across literatures and under different specifications, focusing on entry policies (Mau et al. 2015), repatriation (Şahin-Mencütek and Tsourapas 2023; Cassarino 2020), labor agreements (Kihato and Oliver 2022; Wickramasekara 2015), or issue-linkage and diplomacy (Tsourapas 2017; Thiollet 2011; Betts 2014; Tsourapas 2025; Adepoju, Van Noorloos, and Zoomers 2010). More dramatically, as international migration cooperation develops, often bundling migration with terrorism and organized crime, so have the concerns over implementation. These agreements bolster security forces in non-democratic partner states while lacking the humanitarian protections liberal democratic governments publicly champion. The same governments facing electoral challengers urging them to demonstrate responsiveness on salient issues such as migration (Meguid 2005).

To understand these interrelated transnational and domestic dynamics, I put forth a rationalist approach to international cooperation on migration that bridges government survival concerns and foreign policy design. I combine this with novel data collection to explain the enactment of bilateral *cooperation arrangements on migration* (CAMs) as an interaction among right-wing populist parties and mainstream governments in Western democracies. The data collection advances our understanding by documenting more than 300 agreements across five continents over the last thirty years, directly contributing to the study of international migration cooperation, particularly bilateral security interactions⁴ The data enables us to understand the mode, frequency, and selectivity of agreements previously unknown. For instance, how 67 percent of recorded CAMs are enacted with non-democratic partners and more than a third with closed autocracies.⁵

4. See work by Weiner (1992, [1978]2015), Buzan et al. (1998), Rudolph (2003), Salehyan and Savun (2024), Greenhill (2010), Bermeo and Leblang (2015), Money and Lockhart (2019), Rosenberg (2022), and Lahav and Messina (2024) for the relevance of strategic migration policy in international cooperation and security.

5. Extensive reporting indicates the propensity for these partnerships to raise serious human rights concerns due to the high-likelihood for physical abuse, economic extortion, unmitigated detention, and interpersonal violence in partner states. For example, in June 2021, Libyan coast guards opened fired and almost rammed a small vessel carrying around 50 adults and children in the Mediterranean (Sea-Watch 2021). Two

I propose a theory of *foreign policy co-optation* that explains when and why governments appropriate flexible foreign policy instruments important for the narrative of the opposition. Co-optation and accommodation can keep incumbents in power with limited political cost by reducing electoral challenges (Piven and Cloward 1979; Holdo 2019; Arriola, Devaro, and Meng 2021). I contend that as right-wing populist parties gain seats, mainstream governments engage in foreign policy co-optation, which is the implementation of foreign policy instruments associated with political challengers to reduce the electoral threat those challengers pose to the incumbent. Foreign policy in particular offer government executives a uniquely flexible domain for narrative control. Rather than solely engaging in costly domestic reforms that invite scrutiny and backlash, governments can externalize contentious issues through international instruments which are often negotiated privately and publicly framed as competent responses to crises. Parties can strategically react to right-wing and radical-right challenger parties (Meguid 2005; Abou-Chadi 2016; Van Kersbergen and Krouwel 2008) at the cost of alienating core constituencies or undermining regime legitimacy by shifting their stance (Van Spanje 2010; Hobolt and Tilley 2016; De Vries and Hobolt 2020; Wildi 2024). In this bounded space, CAMs preempt electoral losses by fracturing opposition and tightening narrative control over the issue area (Meguid 2008).

Furthermore, I argue these effects are especially pronounced for governments most challenged by right-wing narratives, rather than ideologically adjoining governments. Studies have aptly recognized the smaller electoral distance in ideal points between conservative and right-wing populist parties, likely helping some parties implement right-wing policies without suffering in upcoming elections (Meguid 2005; Muis and Immerzeel 2017). In contrast, I argue that the ideological distance between mainstream incumbents and right-wing challengers is large, mainstream governments face greater incentives to adopt visible foreign policy signals that demonstrate responsiveness to voter concerns about cultural and

months later, Libyan security forces shot 4 migrants and wounded several more attempting to escape a Libyan detention center (Al Jazeera 2021). Libya receives extensive funding from Italy to control migration from several recorded CAMs.

economic grievances such as migration. This would indicate that left-leaning governments, given their perceived less-restrictive view on migration, would be more challenged to engage with policies that demonstrate control over the issue at hand. Conversely, more conservative governments that are ideologically close to right-wing populist parties are already perceived to be tougher on migration, thus are less challenged by right-wing narratives—to an extent (Williams and Hunger 2022; Hadj Abdou, Bale, and Geddes 2022). CAMs overall offer mainstream governments a politically calculated alternative to more extreme forms of over-reaction or posturing in migration policy without completely dislodging their party brand (Bils 2023; Judd et al. 2017; Green and Jennings 2017).

Given that agreements are often shaped by prior arrangements and broader institutional frameworks, I estimate dyadic gravity-count models of CAMs to account for policy convergence in the frontier state, the partner state, and the historical relationship between right-wing challengers and mainstream governments. My data collection spans ten European frontier states and their partners across Asia, Africa, Europe, South and the Americas from 1990 to 2023. Results support my theory accounting for historical right-wing seat share, colonial relationship, refugee stocks, asylum pressures, and other characteristics.

The project contributes to our understanding of the political right’s impact on international cooperation by providing a generalizable theory that could extend beyond migration to other domains where executives can externalize contentious issues.⁶ It also contributes to existing policy repositioning and party accommodation literature by focusing on tangible foreign policies in contentious issue-specific areas like migration where international constraints and domestic pressures intersect (Meguid 2005, 2008; Abou-Chadi and Krause 2020). The project additionally contributes to migration studies by introducing a novel resource for the quantitative analysis of border externalization and migrant safety with more than 50 individual characteristics cataloging the depth of cooperation.

6. Climate change policy is one such area, as emerging studies show similar dynamics of policy maneuvering to address international climate concerns (Tallent, Jan, and Sattelmayer 2026; Beacham 2023; Genovese 2019).

Lastly, this project sheds light on ongoing policies affecting millions of migrants with little signs of stopping,⁷ that risk legitimizing non-democratic actors at home (Daur 2025) and abroad in an era of deep contestation to the liberal international order (Lake 2026).⁸ This study therefore seeks to support this growing research area by showing how domestic challengers shape democratic governance even when not in power (Chryssogelos 2017; Pevehouse 2020; Haggard and Kaufman 2021; Myrick 2021, 2025; Cottiero et al. 2025; Haggard and Tiede 2025; Carnegie and Clark 2026).

What are CAMs?

Cooperation arrangements on migration are bilateral agreements negotiated to control migration by equating irregular migration to other forms of criminal activity, often solely focusing on irregular migration as a security issue. More specifically, I define CAMs along three dimensions. They are restrictive in direction: their object is to reduce and deter irregular arrivals, not to open, regularize, or protect channels of movement. They are securitized in method: they operate through police, border, military, and intelligence cooperation rather than through labor, development, or humanitarian instruments. Lastly, they are enforcement-oriented in content: they commit signatories to interdiction, surveillance, readmission, and capacity-building for control. An agreement that cooperates on migration without these properties, such as a bilateral labor arrangement, is not a CAM. Establishing this is necessary for the argument that follows, since it is the restrictive content of CAMs that makes them a credible substitute for the challenger’s own policy demands. As Schelling (1980) asserts, and Young (1989) and Milner (1997) point out, cooperation can be *tacit* or

7. Under the recent European Union’s Pact on Migration and Asylum, which entered into application on June 12, 2026, the Pact reiterated the need for close cooperation with third countries at the bilateral level to “promote legal pathways; supporting partners and building their operational capacities in migration, asylum and border management; [and] prevent irregular migration and combat migrant smuggling and trafficking in human beings” (Commission 2026). For specific effects of border externalization, see Alpes (2015), Last et al. (2017), Williams and Mountz (2018), Üstübeci (2019), and Tyszler (2019).

8. For recent system-level analysis discussions on migration, see **money2021globalization** ; Natter and Thiollet (2022) and Fernández-Molina and Tsourapas (2024).

negotiated. CAMs represent negotiated cooperation as they emerge from policy coordination where a bargain is reached. Although I choose to focus on traceable negotiated agreements, many unrecorded arrangements between states on migration and security exist.⁹ However, a key advantage of CAMs is their specificity, which includes records of dates, signatories, concessions, and explicit commitments¹⁰

The collected CAMs are extracted from (1) national online legislative archives, (2) the UN Treaty Collection, (3) national government bulletins, (4) government websites, and (5) for-profit national law databases. CAMs are individually analyzed to verify the meaning of the cooperation and concessions provided. I document CAMs enacted by 10 EU frontier states: Austria, Bulgaria, France, Greece, Hungary, Italy, Poland, Portugal, Romania, and Spain, from 1990 to 2023. The network of CAMs spans 97 countries across Africa, North America, South America, Asia, and Europe, with more than 70 percent of agreements concentrated in Europe, Middle East and North Africa. Figure 1 displays the distribution of CAMs per country over time, showcasing the dissimilarity of enactment across the frontier states.¹¹

CAMs are often negotiated behind closed doors but signed in public. Senior government officials, typically the Minister of the Interior or the Minister of Foreign Affairs, appear alongside their counterparts to formalize the agreement and to emphasize the agreement’s expected benefits for both nations.¹² In the Italian case presented earlier, the Interior and

9. Such is the case of an agreement between Italy and Tunisia in 1998, and Egypt in 2007, where migration and security agreements were reportedly executed but are not available in either country’s online archives (Godenau et al. 2008). Ongoing work will locate these and other agreements.

10. Each CAM is coded across 58 individual characteristics.

11. The CAMs dataset excludes special operations and project-based cooperation, such as the West Sahel Project (2011–2016), the Seahorse Project (2006–2008), and the subsequent Seahorse Network (2007–2008). While these initiatives are substantively relevant to migration management, they constitute operational or technical cooperation frameworks rather than formal bilateral diplomatic agreements enacted by executives (Ippolito, Borzoni, and Casolari 2020). Since the theoretical mechanism of foreign-policy co-optation hinges on visible diplomatic enactment, including such projects would conflate operational enforcement with treaty-based signaling. Moreover, if any of these projects approximate to CAMs either in diplomacy or publicity, a portion of these projects have already been captured in the data. Many of these operations are by-products or related to formal agreements in some form by default. Additional analysis of these operations and projects is a future point in this research agenda.

12. Patterns emerge across the coded arrangements. While countries engage in bilateral relations at different rates and under varying conditions, most agreements link migration to terrorism and organized crime while rarely mentioning humanitarian provisions or migrants’ rights. Partners often reaffirm the value of bilateral friendship, even when they have poor records on civil liberties. A systematic review of these patterns is

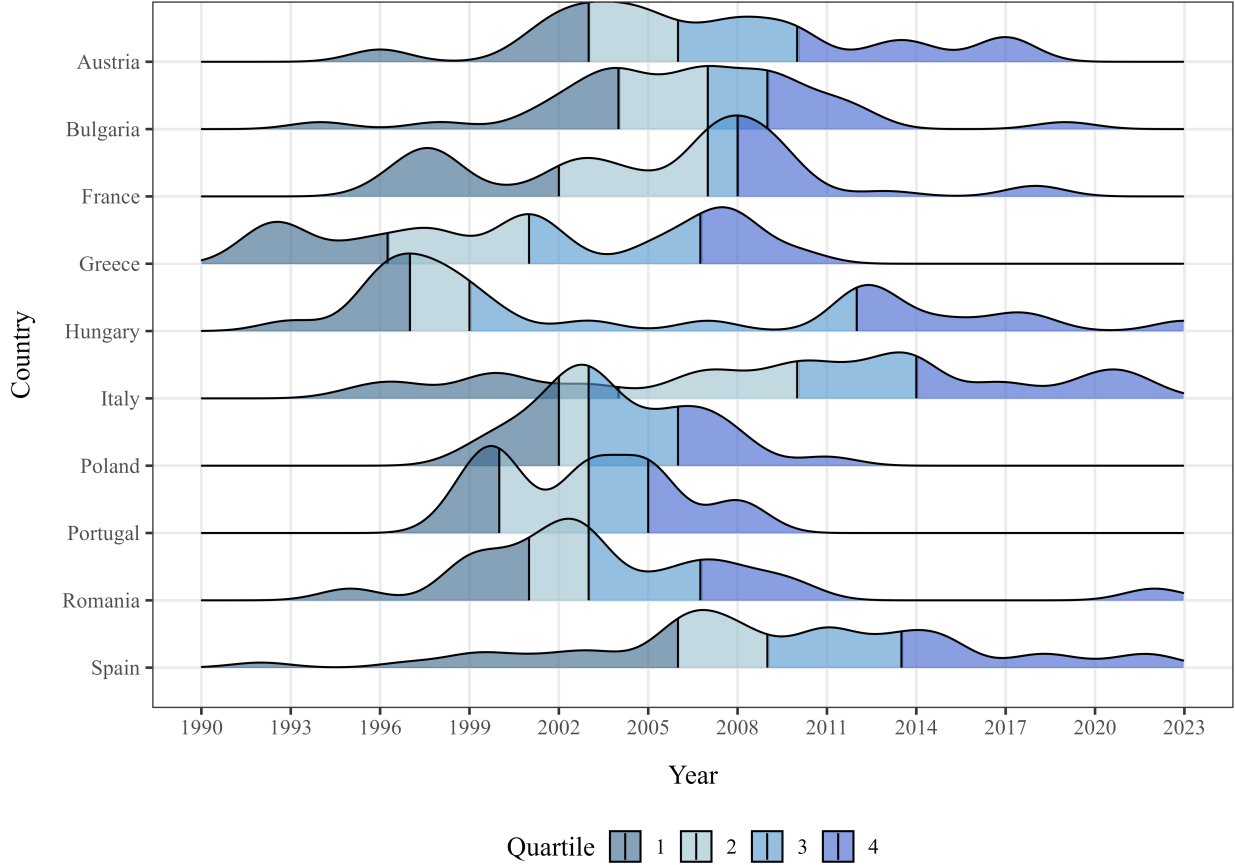


Figure 1: Ridgeline density plot displaying distinct waves of CAM activity across frontier states. The quartile segmentation highlights the temporal clustering of agreements within countries, rather than differences in overall volume.

Ivorian ministers signed a migration and security agreement in a public ceremony that included press coverage and a public reading of the document.¹³ Meanwhile, the UK Minister of State for Security praised ongoing cooperation with Iraqi leadership, noting that “by working together on security, development, and migration, [they] are building stronger relationships that benefit both countries while tackling shared issues such as organized crime and irregular migration” (UK Government, Home Office & FCDO 2025). Such public-facing elements are a common feature of CAMs.

CAMs are assumed to be on average effective in decreasing migration (asylum applications to migrant-recipient countries) but expectations vary. For example, the Treaty of

13. See miltonkwami (2020) for a video recording of the meeting.

Friendship, Partnership and Cooperation (TFPC) between Italy and Libya reduced more than 30,000 asylum applications received in 2008, to only 17,725 in 2009, and barely 10,000 in 2010. The TFPC was instrumental in decreasing migration to Italy until its demise in the Arab Spring in 2011. A more recent example transpired in late 2022 between Spain and Mauritania. The countries reached a bilateral agreement centered on the acquisition of resources and logistical support to combat irregular migration to Spain.¹⁴ Mauritania, where rampant corruption and slavery persist to this day, received technical assistance and recognition as an instrumental partner in stopping migration. While Spain prevented migrants from reaching its mainland and territories. Earlier that year, the joint deployment of Spain's *Guardia Civil* and Mauritania's *Gendarmerie Nationale* intercepted more than 4,000 migrants attempting to leave Mauritania's coast. However, the variation in durability and symbolic salience of CAMs suggests that their effectiveness will not be uniform across all frontier and partner states. Nonetheless, potential policy inefficiencies are unlikely to deter executives from entering CAMs when such agreements contribute to their political survival.

Furthermore, I restrict the collection of CAMs to post-1990 for two reasons. First, analyzing foreign migration policy before 1990 would lend itself to be under the shadow of the United States and the Soviet Union, and their hegemonic relationship to their allied states. For instance, had West Germany adopted punitive measures that categorically excluded migrants from Eastern Europe, such policies would likely have been interpreted as inconsistent with its broader geopolitical standing and would have diminished its reputation. The act of crossing national borders carried strategic weight at the time by showing defiance toward communist governments while simultaneously served the ideological and political interests of Western states (Loescher 2001; Betts 2014, 9).

Second, the Maastricht Treaty (1993) developed the modern EU into its current state by introducing an early attempt of the concept of European citizenship and a shared security and migration policy. Additionally, the 1985 Schengen Agreement (*acquis*) was replaced by

14. Refer to La Moncloa (2022) for the press release published on the official website of the Spanish government and the Council of Ministers.

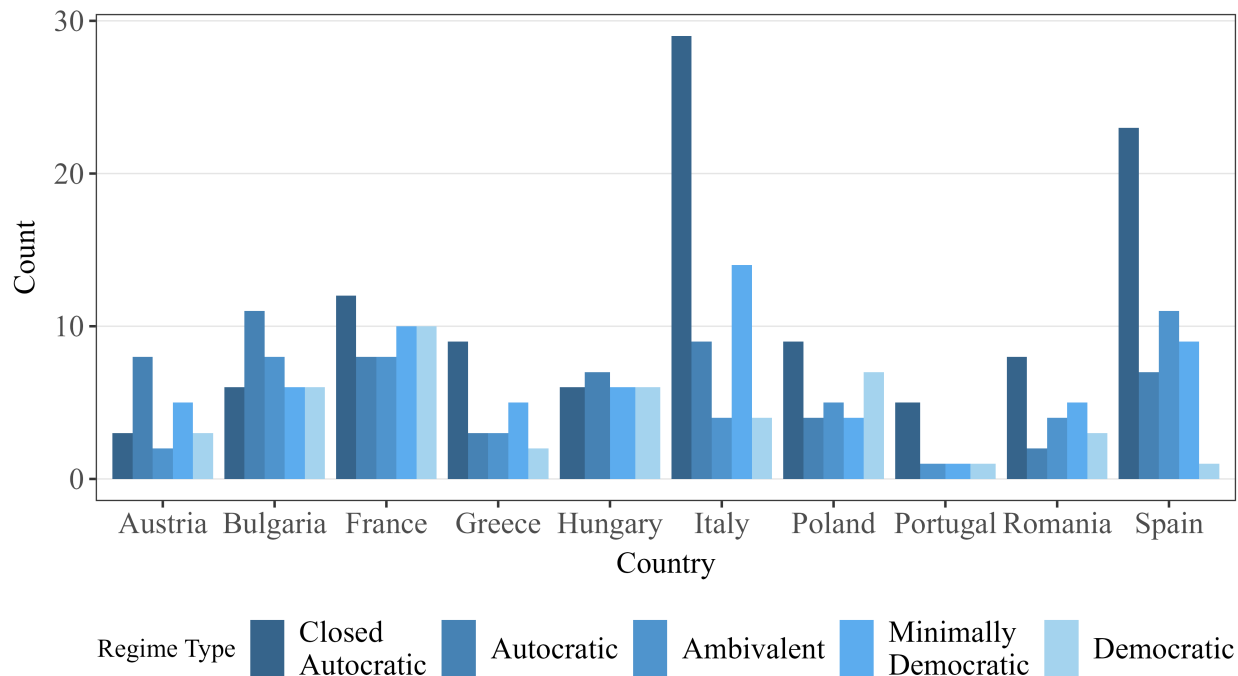


Figure 2: CAMs signed with partners, split by an ordinalized version of the VDEM’s Multiplicative Polyarchy Index (Coppedge et al. 2021).

the 1990 Supplementary Agreement of the Schengen Convention in which Title II brings to the forefront the abolition of internal borders by procuring securitization of its outer borders through cooperation and policing (Huybrechts 2015; Lahav 2004; European Union, n.d., 43). Both events established a stable policy framework that remained largely consistent throughout the period of interest, facilitating the identification of variation across the domestic legislatures with less concern over major systematic changes affecting this study’s analytical scope.

Assuming stability in international hierarchal regimes for the majority of the data collection, and assuming continuity of policies due to the Maastricht Treaty, my next assumption relates to the determinants of bilateralism in a globalized world. CAMs by themselves do not explain how parties electorally benefit from their enactment or co-optation. Explaining this requires connecting the international dynamics of migration cooperation and bilateralism, with the domestic pressures shaping mainstream government internal and external security strategies given their economic and cultural impact (Rudolph 2003).

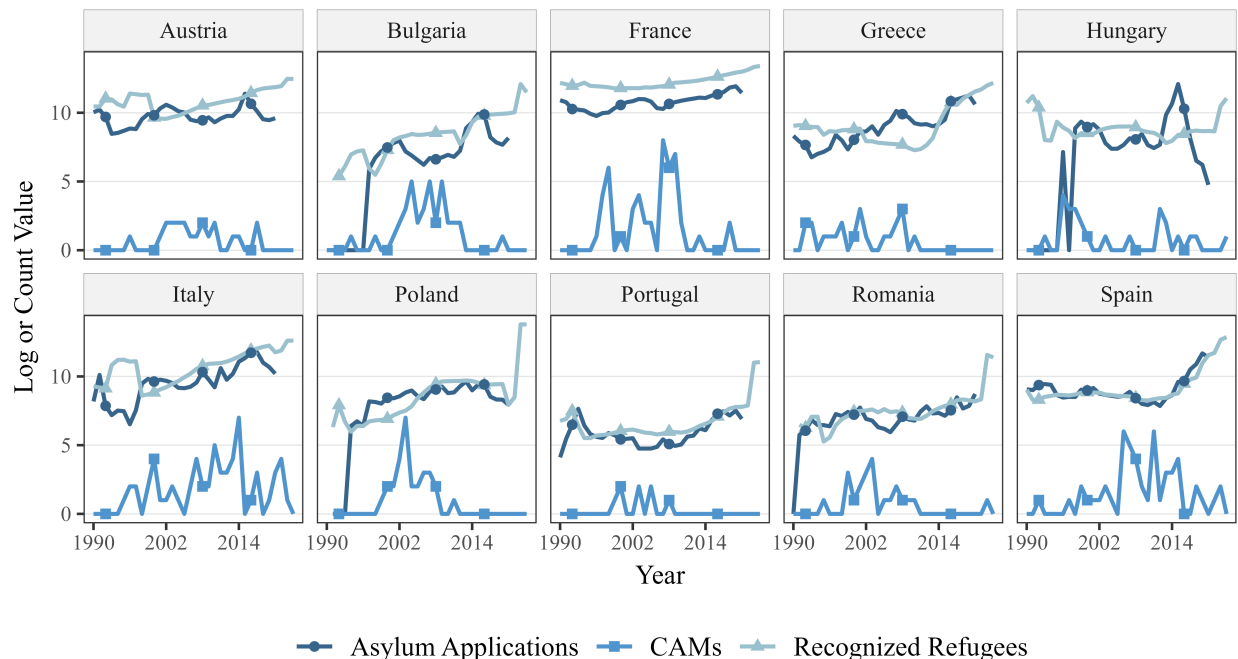


Figure 3: Asylum applications, CAMs, and recognized refugees in the frontier countries.

Bilateralism and Migration Foreign Policy

Bilateralism in migration policy offers solutions to challenges generally found in international cooperation. On the other hand, multilateral migration cooperation amplifies these challenges as it is governed through a fragmented international architecture often referred to as a ‘patchwork’ regime (Geddes 2010). International cooperation emerges from the need to overcome sanctioning, monitoring, distribution, and informational problems (Morrow 1994). Seminal work by Axelrod (1980) and Axelrod and Keohane (1985) additionally illustrates cooperation as an iterative interaction, where reciprocity and future payoffs deter cheating. Reputation is thus a critical long-term constraint on state behavior (Greif, Milgrom, and Weingast 1994; Tomz 2007; Dai, Snidal, and Sampson 2010). Once a state’s reputation is damaged, future cooperation becomes more difficult to sustain as the number of actors involved affects the potential benefits (Olson Jr 1971; Oye 1985; Keohane 1984) but also feasibility (Grieco 1990; Grieco, Powell, and Snidal 1993).

Unlike issue areas governed by formalized global rules, migration cooperation takes place

through dispersed, often informal arrangements involving origin, transit, and destination states. States may view migrants simultaneously as humanitarian cases, security risks, or labor commodities, further complicating preference convergence (Freeman 2006). Some frameworks further highlight migrants' economic value to emphasize their benefits and retrofit economic logic to migration in developed economies (Hollifield 2000; Hollifield, Martin, and Orrenius 2014; Freeman and Ögelman 1998), but the securitization of migration has blunted that narrative for several decades now (Wæver 1995; Buzan 1991; Buzan et al. 1998; Rudolph 2003; Greenhill 2010). While converging on migration agreements helps overcome collective action failure on a humanitarian problem with large spillover effects (Tsourapas 2017; Betts 2014, 62), this is not always successful. This lack of overarching norms or enforcement mechanisms incentivizes governments to pursue selective cooperation where they can maximize control while avoiding the costs of poorly-planned binding commitments. The changes brought by globalization, given that migration is an economic and sociological, domestic and international issue, makes cooperation on migration especially suited for bilateral attention.

Bilateral agreements simplify cooperation by matching benefits and concessions with a single partner. When an issue is salient and contested domestically, international agreements give executives a way to act visibly on the issue while distributing the political costs internationally. This allows for selective cooperation, underpinned by domestic pressures and facilitated by historical ties, geographic proximity, or mutual strategic interest (Money and Lockhart 2019; Kinne 2013). For example, Italy's enduring cooperation with Libya reflects both colonial legacies and geographic calculations. In the EU context, the first article of the Maastricht Treaty exemplifies that even under a shared union "decisions are taken as closely as possible to the citizen" (Lahav 2004, 156). Italy can better tailor migration controls to each partner's capacity, preserving its policy preferences while minimizing collective bargaining costs and still influencing the issue at hand with lesser EU or additional partner involvement.

It is true that transnational interdependence has complicated policy development, bilat-

eral or otherwise (Hollifield 1992; Del Sarto 2021). There will always be norm entrepreneurs that advocate for their preferred policies at the domestic level drawing from international norms or ideologies (Hobolt and Tilley 2016; De Vries and Hobolt 2020). Yet, although domestic policy development is irrevocably intermingled with international constraints (Gourevitch 1978; Putnam 1988; Evans et al. 1993), a mediating factor will still be the governments’ desire to adhere to their electorate’s demands. Whether we assume supranational entities force the enactment of particular policies or simply encouraging normative changes, international influence will still have to endure domestic institutions and their own processes of representation (Risse-Kappen 1995; Finnemore and Sikkink 1998; Checkel 2001; Cheong and Haggard 2023).¹⁵ This domestic filtering helps explain why governments will still use bilateral migration policies due to their strategic flexibility and political ownership, even if internationalism and multilateralism remain constant in the policy maker’s toolkit.

Taken together, bilateralism offers executives autonomy, speed, and targeted cooperation with limited transparency and accountability. CAMs are thus functional and politically efficient. For governments under pressure from rising challengers, bilateralism generally provides a low-cost, higher-control mechanism to signal enforcement capacity without multilateral delays. CAMs specifically reflect not only efficiency-seeking behavior but also strategic asymmetry: frontier states externalize border enforcement while maintaining discretion, narrative control, and electoral signaling capacity.

Mainstream Governments and Challenger Parties

Of course, international cooperation requires vertical accountability of domestic dynamics—however contentious these might be. Recent studies have shown the success of right-wing populist parties in Western democracies driving mainstream governments to adjust their policies and underscore their usefulness in government (Wildi 2024; Williams and Zeager

15. More on multilateralism (Ruggie 1982, 1992; Ruggie et al. 1992; Levy, Young, and Zürn 1995), international policies changing domestic values in Germany (Checkel 2001), and policy influence and its mediation (Risse-Kappen 1995; Biersteker et al. 2009)

2004; Lewandowsky and Wagner 2023; Holdo 2019; Piven and Cloward 1979). The 1970s saw the rise of ‘challenger’ parties as post-materialist values spurred the formation of new political movements. These parties aimed to expand the range of democratic engagement and supplement dominant mainstream parties such as conservative, liberal, Christian democratic, and socialists, laying the groundwork for the ‘counter revolution’ of right-wing parties in more recent years (De Vries and Hobolt 2020; Inglehart 1977; Lipset, Lipset, and Rokkan 1967).

On the other hand, co-optation can reduce the organizational strength of the opposition, but the methods employed vary across regimes and desired outcomes (Arriola, Devaro, and Meng 2021). In less pluralistic systems, governments often rely on tools such as repression, patronage, and information control to maintain power (Gehlbach and Sonin 2014; Hollyer, Rosendorff, and Vreeland 2015; Tyson and Smith 2018). Democratic governments bounded by established norms and institutional restrictions instead turn to policy co-optation as a strategic alternative. Governments are aware of the volatility that immigration politics can bring and seek to decrease the political influence that right-wing populist parties might gain by claiming the incumbent as unfit to address migration. In the case of migration, these pressures can be further exacerbated by highlighting security concerns such as terrorism (Helbling and Meierrieks 2022) and concerns over labor (Scheve and Slaughter 2001). Rising asylum applications, recognized refugee numbers, and highly visible migration-related events ensure that immigration remains a salient and electorally charged issue (see Figure 3).

The current era of voter volatility, characterized by the rise of single-issue movements and challenger parties, highlights this transformation of the political landscape (Poguntke 1987; Meguid 2005; Lahav 2004; De Vries and Hobolt 2020). But while challengers are relatively easy to identify as ideological or organizational outsiders, mainstream parties and governments can be slightly harder to define (Moffitt 2022).

I define mainstream parties as those affiliated with moderate or historically dominant ideological traditions within a given political system closely echoing the notion of orthodox

parties discussed by Pop-Eleches (2010) and Moffitt (2022). These orthodox parties include ecological or green, socialist, social democratic, liberal and Christian democratic, and conservative parties (Lehmann et al. 2024). A mainstream government, in turn, is one composed primarily of these parties and led by a prime minister (PM) from this ideological family. Fortunato and Adams (2015) find that voters “map the PM party’s policy position onto the junior coalition partners” while little evidence for a reciprocal effect can be found. Subsequent work by Fortunato et al. (2021) also finds a consistent ordering of ‘cue weights’ for government roles placing the PM at the top under coalition governance. Thus, it is essential to conceptualize mainstream government not only by coalition composition but also by the PM’s partisan affiliation. For instance, a coalition with a few cabinet posts held by challenger parties may still be perceived as “leftist” if the PM is a social democrat, prompting continued resistance from right-wing populist voters.¹⁶ Yet, for mainstream governments to successfully co-opt the narrative and bring back voters to their side—if at all possible—they must co-opt policies that resonate with right-wing populist voters and their preferences.

The Political Right and CAMs

Although often labeled as ideological and organizational outsiders, I explicitly define right-wing populist parties as those who seek to crystallize the demands of the in-group using a combination of anti-establishment, people-centric, and anti-elitist remarks, with often xenophobic undertones when describing migration.¹⁷ Migration sits at the center of these claims by combining exclusionary remarks and disdain for globalization, often incorporating public smear campaigns against immigrants and fear mongering.¹⁸ Right-wing parties are assumed to be unitary actors with a combination of policy-seeking and vote-seeking models

16. Additional mainstream government specification is found in the *Empirical Approach* section.

17. Right-wing populist parties focus on the normative distinction between the ‘elites’ and the ‘people’, criticizing political correctness, and expressing divisive views of the world (Taggart 2000; Mudde 2007; Rooduijn et al. 2017; Shehaj, Shin, and Inglehart 2021). Another popular definition: populism as a ‘Manichean discourse’ that separates good people versus evil elites.

18. These fears might not be unfounded as recent events could have primed people to feel uneasy about migration or particular migrants, yet these are expanded and contrived to further increase angst in the population (Williams and Hunger 2022).

of party behavior, following Strøm (1990) and Müller and Strøm (1999).¹⁹

Policy-seeking parties seek to maximize their impact on policy and would align closely to parties spatially adjacent to their position to some extent, and can differentiate between beneficial policy outcomes from a range of options (Müller and Strøm 1999). Enacting policies deemed of high intrinsic value would therefore constitute an electoral win (Lubbers, Gijsberts, and Scheepers 2002; Rydgren 2008; Norris and Inglehart 2019). On the other hand, vote-seeking models assume parties to be vote-maximizers to maintain and increase their influence in parliament (Downs 1957; Müller and Strøm 1999). Since right-wing populist parties know they are chosen for their promises in a broader ideological debate as much as their party platforms, they can safely assume most policies that address a core tenet of the party will be policy gains and in turn increase their vote share.

Under these conditions, CAMs appear as a strong combination of right-wing populist parties' core tenets and could thus provide electoral benefits if enacted by them. CAMs are expected to at least exert a nominally larger control over migration (asylum applications and recognized refugees), establish security cooperation channels to monitor corridors,²⁰ and signal stark measures to fulfill their promises compared mainstream governments. Other restrictive instruments, such as region-specific entry bans or comprehensive border closures, are available to governments but are institutionally constrained or fiscally costly. Two further features make them less attractive than CAMs for co-optation. Their costs fall on identifiable domestic constituencies, such as employers, universities, diaspora communities, and border regions, which gives those groups both the motive and the standing to mobilize against them. This means their implementation is comparatively easy for voters and the press to verify, so a government that adopts them can be held to the observable results. CAMs invert both

19. Scholars should not ignore the importance of particular figures driving major intra-party politics, such as Matteo Salvini from Italy's *Lega Nord* or Northern League (LN), or the Kaczyński brothers in Poland. Recent and upcoming work expands upon the role of personalist parties in changing democracies (Frantz et al. 2021; Frantz et al. 2022).

20. 'Corridors' refer to routes taken by migrants and smugglers to reach a destination country such as the West African, Balkan or East Mediterranean routes that reach Europe, and the Darién Gap used to reach the US.

properties given that the costs are borne abroad and by migrants who are not part of the domestic electorate, and the implementation is remote enough that the signing itself carries the signal. Lastly, multilateral approaches, particularly those involving funding allocations to international organizations, may conflict with the anti-multilateral, anti-elitist, and isolationist orientations of right-wing populist parties. CAMs occupy a functional middle ground between border closures and mass deportations as right-wing populist parties are tethered to popular opinion and reputational costs just as any other party (Money and Lockhart 2019; Akkerman, De Lange, and Rooduijn 2016, 19–20).²¹ While CAMs carry potential reputational costs for governments engaging with illiberal partners (Weisiger and Yarhi-Milo 2015; Hafner-Burton 2008), their domestic political benefits in managing voter demands often outweigh these risks, especially as governments are dynamic in policy their policy responses when competing on migration (Gessler and Hunger 2022; Green and Jennings 2017). Nonetheless, the strategic environment must be conducive for co-optation.

A Theory of Foreign Policy Co-optation

The likelihood of foreign policy co-optation depends on two key contextual conditions: a significant presence of the opposition in the legislature and the political salience of the issue at hand. In this case, this translates to a credible right-wing populist presence in parliament and a politically salient threat from migration. I posit right-wing populist parties' narrative dominance over migration drives mainstream incumbents only after that narrative is converted into parliamentary seat share. While popular mobilization matters, it is sustained electoral support that poses a real threat to re-election. Furthermore, office-seeking governments are most vulnerable when the issue area is salient for voters (Bils and Izzo 2025; Bils 2023; Holdo 2019; Piven and Cloward 1979, 30). Co-optation is most likely when migration remains an actively contested narrative. While issue salience may fluctuate over time due to voter memory and institutional dynamics, governments face acute pressure to act once these

21. A deeper exploration of CAMs and how they compare to other policies is presented in subsequent work.

conditions are present.

The government then faces a binary choice: implement CAMs as a low-cost, administratively tractable policy instrument, or maintain the status quo. In the absence of electoral pressure, governments would prefer to manage CAMs and broader migration policy on their own timeline. However, when both conditions are satisfied, credible right-wing populist seat share and high issue salience, the decision horizon compresses, incentivizing rapid policy action. CAMs are particularly well-suited to this constraint, given that their political signaling value is realized at the moment of signing rather than enforcement.²² The electorate will punish or reward governments over CAMs enactment reflected in right-wing populist seat share. Governments understand that the choice of voters to punish relies on their preferences and beliefs, which might not just be informed by the policy itself. Policy diffusion, ideological voting, exogenous shocks, and misinformation could undermine co-optation and in turn maintain right-wing populist voters in the long term (Van Kessel, Sajuria, and Van Hauwaert 2020; Ahlquist, Copelovitch, and Walter 2020). Nonetheless, since the government is seeking to remain in office and migration narratives are being controlled by the challenger, they must act accordingly to signal competence.

Governments, as the main executors of CAMs, strategically time their implementation to maximize public attention and political gain. Yet tracing exactly *when* CAMs will be enacted becomes difficult in the context of immigration. Trends and endogenous shocks would maintain migration a constant focus regardless of real-time numbers.²³ Governments are therefore compelled to respond to the political salience of migration rather than to its empirical trends.²⁴ A recent American example illustrates this logic. During the 2024 campaign, President Biden met President Trump at the border to blunt Republican attacks on

22. Refer to Appendix A.5. According to the data, two-thirds of CAMs are enacted within the first 2 years after signing, though many of these informally begin preparations or action much before (Ippolito, Borzoni, and Casolari 2020, 10).

23. This might also be further complicated by the deterioration of race relations between groups and the assignment of racial tensions to migrant flows present in European society.

24. This does not imply that CAMs can never align with actual increases in migration pressure, only that such alignment is not required given the primacy of migration's political salience.

migration by visibly engaging the issue, going as far as asking Republicans to support the “toughest...most effective border security bill this country’s ever seen” (Gutierrez, Marquez, and Lebowitz 2024). The perceived weakness of Biden’s policies was heightened by Republicans at the time, although in reality April 2024 saw 30 percent lower border encounters in the southwest border than April 2023.²⁵ It must be noted that the urgency of this signaling is not merely rhetorical. Governments that fail to respond to right-wing populist pressure face a compounding threat: as challenger parties accumulate seat share, parliamentary fragmentation accelerates, shortening government survival and narrowing the window for policy action. Mainstream governments across Europe have repeatedly confronted this dynamic for some time now. For example, the rise of the Freedom Party of Austria (FPÖ) in the early 2000s caused almost all Austrian parties to adopt more stringent political views on migration and asylum seeking (Duncan 2010; Albertazzi and Mueller 2013).²⁶

Similarly, Denmark’s leftist party Social Democrats (SA) have adopted right-wing stances on migration (Khalid and Mortensen 2022). The SA revealed their new party policy in 2017 demonstrating a right-wing stance on migration and essentially “outflank their competition” with restrictive immigration bills (O’Leary 2019). Their election success in 2019 came as a result of co-opting right-wing policies to appear stronger and more fit to lead (Khalid and Mortensen 2022). This follows from a decrease in interest in accepting asylum seekers, and the passing of legislative bill L226 which would involve third-country processing and resettlement outside of the EU (Lemberg-Pedersen, Whyte, and Chemlali 2021). In these instances, governments enacted policies to remain in office and attempted to appropriate the issue at hand.

Furthermore, foreign policy co-optation can mitigate deeper institutional threats. Co-optation helps mainstream governments counter parliamentary fragmentation and safeguard

25. Refer to U.S. Customs and Border Protection (2024) for US Customs and Border Protection report.

26. In 2002, leadership of Spain’s People’s Party (PP) explicitly linked stringent migration policy to the goal of preempting a domestic surge in far-right voting, citing far-right voting trends in neighboring France (Cornelius Wayne et al. 2004, 392). The People’s Party went as far as rolling back migrants’ rights to public services in the new version of the *Ley de Extranjeria* first implemented in January 2001 (406).

incumbency. Right-wing populist parties are an additional option to voters shopping for representation, gradually increasing turnout rates as feelings of confidence in representation increase (De Vries and Hobolt 2020, 212). This increase of choices unfortunately creates fragmentation within legislatures. The more parties are involved in governance, the more difficult it becomes to please them all, and the harder it is to stay in government steadily (Sani and Sartori 1978; De Vries and Hobolt 2020, 223). Governments with a higher vote share for challenger parties (right and left wing) break down at an increasing rate right after 400 days, compared to the average government that breaks down at the same rate past the first 600 days (De Vries and Hobolt 2020, 226). In this context, CAMs provide a low-cost tool to counteract parliamentary fragmentation in future elections as the Spanish, Austrian, Danish, Italian, and American examples help illustrate. This brings me to my first conjecture:

Hypothesis 1

Mainstream governments are expected to enact more CAMs as the seat share of right-wing populist parties in parliament increases.

My second conjecture relates to the ideological family of the government. Governments most challenged by right-wing populist narratives, rather than ideologically adjoining governments, will face stronger pressure to enact CAMs to remain in power. When the ideological distance between mainstream incumbents and right-wing populist challengers is large, mainstream parties face greater incentives to adopt visible foreign policy signals that demonstrate responsiveness to voter concerns. But in doing so, they can still incur electoral costs from their own parties. Ideologically distant governments such as ecological or green, socialist, and social democratic parties, risk losing the most support if they implement policies beyond their ideological position, thereby alienating their loyal electorate while still being punished by the voters they aim to appease. Conversely, more conservative governments that are ideologically close to right-wing populist parties are already perceived to be tougher on migration thus

less challenged by right-wing populist narratives, yet still susceptible to narrative challenges (Williams and Hunger 2022; Hadj Abdou, Bale, and Geddes 2022). Whether conservative governments actually enact stronger migration policies is inconsequential to the narratives being pushed by right-wing challengers *and* to the perception of the electorate, as long as the narrative is controlled by right-wing challengers. Though I expect any government to enact CAMs to combat right-wing populist opposition, my second conjecture states that:

Hypothesis 2

Left-of-center governments are expected to enact more CAMs as the seat share of right-wing populist parties in parliament increases.

My third conjecture relates to the regime of the partner country. A government's co-optation strategy may come at the expense of its international reputation by binding it to partners with poor humanitarian records. Scholars of international human rights enforcement have demonstrated that international audiences, including NGOs, international organizations, and foreign governments, impose real reputational costs on states that cooperate with illiberal actors (Hafner-Burton 2008). The conventional expectation drawn from this literature is that democratic governments should be reluctant to enter cooperation agreements with non-democratic partners, particularly on sensitive issues involving the movement and treatment of vulnerable populations. However, cooperation with democratic partner states dealing with similar migration pressures is less consequential to this calculation, as their involvement carries minimal reputational exposure and likely less normative costs. I assume that democratic and non-democratic alike generally prefer the ability to cooperate with their counterpart on a politically salient issue, likely equally affecting the partner state, as they reap material benefits such as training, capability building, funding, intelligence, and communication. Nonetheless, it is cooperation with non-democratic regimes that can raise costly strategic and normative dilemmas if the material benefits it receives ultimately undermine migrant safety, the safety of the population within the partner state,

or altogether threaten the nature of the relationship between frontier and partner state.

CAMs may foster codependent bilateral relationships that may yield additional strategic leverage. The non-democratic partner’s willingness to engage in CAMs can vary depending on potential benefits—such as funding, training, recognition, and access to restricted information—which they weigh through their own cost-benefit calculations (De Mesquita et al. 2005; Svolik 2008, 2012). Third-country partners, aware of Europe’s difficulty managing migration, can exploit this dependence by threatening to release migrants as a coercive tool, as both Gaddafi, Erdoğan, and Lukashenko²⁷ managed to do shortly after cooperating with the EU and selected frontier states. As Greenhill (2010) documents, weak states can relatively effectively weaponize migration by threatening to release detained migrants, even beyond the European context. These could result in migrant abuse, further halting for good-governance goals (Weisiger and Yarhi-Milo 2015) and providing reasons to be named and shamed (Hafner-Burton 2008; Jervis 1978, 169), especially if they flaunt their agreements similar to Berlusconi working with Gaddafi or the British government with Rwanda (Bullen and Bartram 2022). Less-humanitarian partner governments can additionally benefit from retaining migrants by exploiting them as cheap labor (Shin 2017; Miller and Peters 2020), or from insulating their populations from democratic ideas and institutions (**petersmiller2020restraining**). As Weiner (1992) and Zolberg (2006) note, migration has been a foremost nation-building tool for centuries. Governments strategically determine who to expel, who to contain, and who to let leave to better fit their vision.²⁸

The foreign state can exploit that codependency by stockpiling migrants and threat-

27. Lukashenko orchestrated a migrant crisis starting in August 2021 by facilitating flights from the Middle East and North Africa region to Belarus, then directing migrants to breach the Polish border, with Lithuania and Latvia as well as Poland declaring states of emergency shortly after. Though Lukashenko’s attempt to get back at the EU’s lack of recognition of his government and subsequent sanctioning, success was thwarted by further sanctioning. Nonetheless, migrants used as pawns suffered extensively, with at least 20 migrants dying due to frigid winter conditions, and many others across Belarus, Poland, Latvia, and Lithuania (Gera and Grieshaber 2022).

28. Even while cooperating, states may selectively restrict emigration to exploit migrant labor or remittance flows (Andresen, Bergdolt, and Margenfeld 2012; Leblang 2017), or allow dissidents to leave (Freeman and Ögelman 1998). Making their cooperation even more dynamic than expected and inviting further research into this interaction.

ening their release in exchange for additional benefits (Betts and Collier 2017; Project 2017; FitzGerald 2020). In this sense, CAMs can create strategically useful codependent relationships that might be normatively costly for democracies. Yet, mainstream governments remain primarily motivated by electoral incentives regardless of the partner state’s liberal democratic values. These reputational costs are undermined by combating electoral challengers, addressing legislative fragmentation, and managing migration. Foreign policy relating to migration can additionally positively combat feelings of economic decline and insecurity regardless of party alignment (Buzan et al. 1998; Huysmans 2000; Menjívar 2014), further decreasing the costs to making particular policy actions. The normative costs associated with working and supporting migration management in non-democracies are set aside by governments rational approach to right-wing populist challengers. This brings me to my third conjecture:

Hypothesis 3

The regime of the partner country will not significantly affect the propensity to implement CAMs.

It must be noted that though this study acknowledges the dangers of enacting particular foreign policy, this condition might not be present in other applications. The theory as it is currently presented simply defines the current parameters but not its limitations. The foreign policy co-optation mechanism presented operates in (1) a competitive electoral system in which challenger parties can credibly threaten the survival of the government through parliamentary seat share, (2) where a mainstream government with executive discretion over foreign policy is aiming to stay in power, and (3) when a salient issue area’s political narrative is dominated by the challenger. These conditions are broadly satisfied across Western parliamentary democracies and likely extend to non-Western parliamentary systems where political challengers or entrepreneurs and executive foreign policy flexibility are both

present (De Vries and Hobolt 2020). This might allow for a higher rate or deeper engagement with foreign policy decisions suitable for government survival. Subsequent work is imperative to understand its application in similar and dissimilar context. This study initiates this analysis with set of frontier states serving as the first point of contact for migrants, where these dynamics are present for its analysis.

Empirical Approach

My objective is to explain the incidence of agreements by mainstream governments, measured as annual counts of agreements signed by 10 frontier states (Austria, Bulgaria, France, Greece, Hungary, Italy, Poland, Portugal, Romania, Spain) while accounting for frontier-specific and dyadic dynamics that might affect their signing. I employ a Poisson Pseudo-Maximum Likelihood (PPML) as the main estimator with high-dimensional fixed effects to model the annual count of CAMs (Leblang and Helms 2023). The PPML estimator is appropriate given the count construction of the dependent variable, the presence of excess zeros, and potential heteroskedasticity in dyadic data (Gourieroux, Monfort, and Trognon 1984). Unlike OLS, the PPML estimator remains consistent under heteroskedasticity (Silva and Tenreyro 2006), behaves well in dyadic environments (Weidner and Zylkin 2021), and performs reliably even when the data contain excess zeros or suffer from overdispersion (Silva and Tenreyro 2011).²⁹ I include origin-country fixed effects (θ_i) and decade fixed effects ($\tau_{d(t)}$) to absorb time-invariant heterogeneity across origin states and decade-level shocks, a structure for which fixed-effects PPML performs reliably (Weidner and Zylkin 2021). The predictors (β, δ) and controls (γ) are semi-elasticities which allows for straightforward interpretation of coefficients as well (Silva and Tenreyro 2011).

I therefore estimate the equation, letting y_{ipt} denote the count of CAMs from origin i to

29. The natural alternative of log-transforming the outcome and estimating two-way fixed-effects OLS can be problematic for this set up given that my dependent variable is zero in the overwhelming majority of dyad-years, so $\log(y)$ is undefined for most of the sample. PPML accommodates the zeros directly while retaining the multiplicative interpretation; but that said, linear models estimated on the country-year collapse are reported in Appendix A.7 as a robustness check.

partner p in year t :

$$y_{ipt} = \exp \left(\beta_0 + \sum_{j=2}^4 \beta_j \mathbf{1}\{G_{ipt} = j\} + \beta_R RWP_{ipt} + \sum_{j=2}^4 \delta_j \mathbf{1}\{G_{ipt} = j\} \cdot RWP_{ipt} + \mathbf{X}'_{ipt} \boldsymbol{\gamma} + \theta_i + \tau_{d(t)} \right) + \varepsilon_{ipt} \quad (1)$$

Where G_{ipt} is the type of government in origin i in year t , and $\mathbf{1}\{G_{ipt} = j\}$ is an indicator for each of the three non-reference types ($j = 2, 3, 4$: *Socialist*, *Social Democratic*, and *Conservative*), with *Center Right* as the reference category, and $\mathbf{X}_{ipt}$ is a vector of dyadic, partner-state, and origin-state controls with coefficient vector $\boldsymbol{\gamma}$. Government type is measured using either the *PM* or the *PM + Cabinet* classification, and I estimate the model separately for each. RWP_{ipt} is the seat share of right-wing populist parties. The coefficients of interest are the interactions δ_j , which capture how the association between RWP seat share and CAM signing differs by government type relative to Center Right. Standard errors are clustered by origin country.

The dataset comprises 9,458 directed dyad-year observations spanning 10 frontier states and 97 partner countries from 1990 to 2023, yielding 276 unique dyads. The 97 partners are the union of partners observed across all frontier states as no frontier state negotiates with all 97. In relation to CAMs, Social Democratic governments account for the largest share of CAMs (37.3 percent, $n = 122$), followed by Conservative (25.7 percent, $n = 84$) and Center Right governments (23.9 percent, $n = 78$). Right-wing populist governments account for a modest share of CAMs (6.1 percent, $n = 20$), but are removed for the analysis. Socialist governments account for a smaller but theoretically important share (5.8 percent, $n = 19$), while Special governments are rare (1.2 percent, $n = 4$). Notably, the mean CAM rate per dyad-year is broadly similar across Social Democratic (0.032), Conservative (0.042), Center Right (0.036), Socialist (0.034), and Right-wing populist (0.028) governments. Complementary OLS models are reported in Appendix A.7.³⁰

30. Results for PPML are additionally crosschecked using the `ppmlhdfc` package in Stata as per Correia,

Main Explanatory Variables

Despite its centrality to the argument, mainstream government resists a one-size-fits-all definition given the considerable variation in party systems and ideological landscapes across countries and time. I identify mainstream governance by moderate or historically dominant ideological traditions within post-Cold War European party systems, and I construct two measures that differ in their treatment of coalition composition, which allows me to assess robustness across specifications. The first, *PM*, classifies each government by the party family of the prime minister. The second, *PM + Cabinet*, combines the PM’s party family with the partisan composition of the cabinet in a rule-based assignment. Although CAMs are formally signed by ministers, typically of the Interior or Foreign Affairs, the decision to pursue them often involves the prime minister, as with Meloni’s engagement with Albania and Berlusconi’s with Libya. The PM’s prominent public visibility and agenda-setting authority, particularly under majoritarian or semi-presidential systems, make the PM’s party the relevant locus of political agency. Voters also map government policy positions onto the prime minister’s party, making the PM’s affiliation the most policy-relevant signal for electoral accountability (Fortunato and Adams 2015; Fortunato et al. 2021).

Party family classifications are drawn from the Manifesto Project (Lehmann et al. 2024) and the Party Government in Europe Database (PAGED) (Hellström et al. 2024a), cross-referenced and manually coded where parties are unclassified or ambiguous. Party families include Green or Ecological, Socialist, Social Democratic, Liberal, Christian, Conservative, and Right-Wing Populist or Radical-Right parties. I collapse these into five mutually exclusive government types: Socialist (including Green or Ecological parties), Social Democratic, Center Right (including Liberal and Christian parties), Conservative, and Special.³¹ I exclude Special governments due to low n , and Right-Wing Populist governments (including Radical-Right) because these parties are the challengers to mainstream governance.

Guimarães, and Zylkin (2019).

31. Special is a residual category for agrarian, ethnic, and populist catch-all parties that do not fit cleanly into the main families.

For *PM*, each government is assigned a type based solely on the PM’s party family. For *PM + Cabinet*, governments are classified using a weighted combination of the PM’s party family and the partisan composition of the cabinet, counting the PM’s party twice to reflect the prime minister’s disproportionate influence over the government’s perceived ideological position. Governments are coded by rule-based thresholds on the proportion of cabinet parties on each side of the ideological spectrum. For example, a cabinet led by a Social Democratic PM and composed of at least 50 percent left-leaning parties (Socialists, Social Democrats) is coded as Social Democratic. Figure 4 shows how classifications vary across the two measures, illustrating cases where coalition composition shifts the coding relative to the PM’s party family alone. A limited number of cases (e.g., Italy, Poland, Bulgaria) required ex-post adjustment where the source datasets masked ideological alignment or lacked a family assignment, based on secondary-source triangulation and the Manifesto Project and PAGED codebooks. See Appendix A.6.

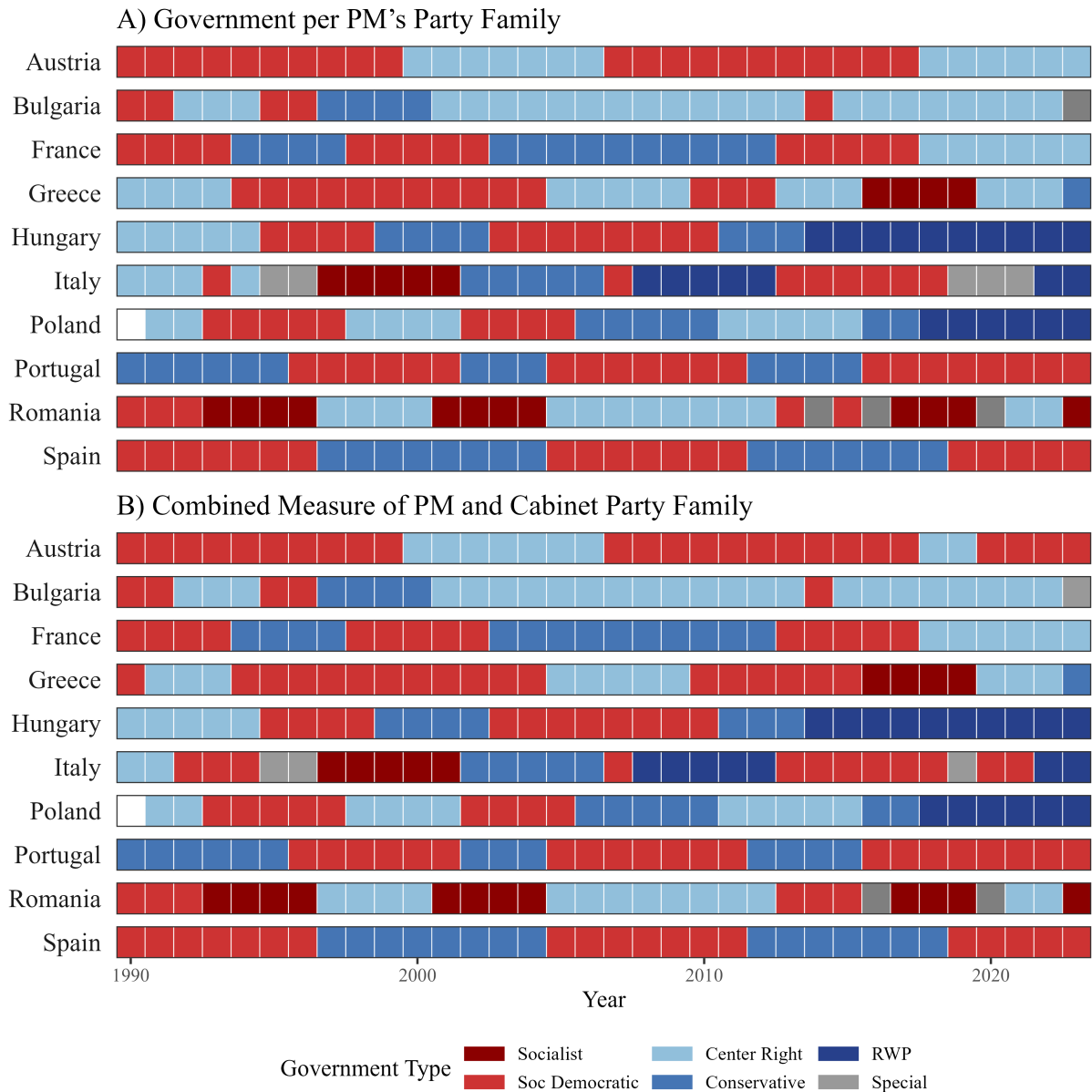


Figure 4: Change in government party families over time by country. Panel A shows the PM's party family. Panel B shows the combined measures, including PM and cabinet family. Tile color indicates party family and government type from 1990 to 2023.

Right-wing populist seat share (*RWP Seat Share*) measures the combined seat share of parties labeled “Right-Wing Populist and Radical-Right” following Mudde (2007) and Carter (2013). Parties and coalitions absent from the Manifesto Project were manually coded following Lehmann and Zobel (2018), with secondary sources used for Bulgaria, France,

Greece, Hungary, Italy, Poland, and Romania. A full list of manually coded parties is in Appendix A.6.

Dependent Variable

The count of CAMs per dyad-year acts as the dependent variable for subsequent results. The CAMs dataset contains 327 collected arrangements from January 1, 1990 - December 31, 2023 (see Appendix A). 58.5 percent of CAMs were enacted with a European partner, 18.2 percent with partners in the Middle East and North Africa, with the remainder distributed across Africa, Asia, and the Americas.³² A majority of dyad-years involve non-democratic partners with 66.8 percent ($\text{MPI} \leq 0.6$), of which a 34.5 percent are closed autocracies ($\text{MPI} \leq 0.2$).³³ Among CAMs signed with non-democratic partners, 46.1 percent are enacted with non-democratic European partners, 26.3 percent with non-democratic Middle Eastern and North African countries, and 11.5 percent with Asian and Pacific partners. Non-democratic partners in West and Central Africa account for 9.2 percent of all non-democratic partners. Figure 5 shows this distribution with Italy, France, and Bulgaria accounting for the largest shares of CAMs signed with European partners. I additionally lead the dependent variable by $t+1$ and $t+2$ to observe delayed effects of government type and right-wing populist strength on CAMs formation.

I control for several predictors of CAM formation, including refugee stock and asylum applications in the frontier state, drawn from The World Bank (2024) and Eurostat (2024) respectively. Asylum applications are measured as the annual count of first-time applicants from each partner country to each frontier state, while refugee stock captures the total number of refugees hosted by the frontier state. Both variables operationalize migration pressure,

32. Region names designated by the UNHCR sub-region coding.

33. I denominate non-democratic any country that is identified as *Ambivalent*, *Autocratic*, or *Closed Autocratic* at the time of the CAM given their Multiplicative Polyarchy Index (MPI). This translates to an operationalization of less than 0.60 out of 1 (Coppedge et al. 2021). Conversely, democratic states are those defined as *minimally democratic*, or *democratic*, based on the ordinalized version of the MPI measure. Figure 2 shows the distribution of the partner countries given their regime type (Coppedge et al. 2021).

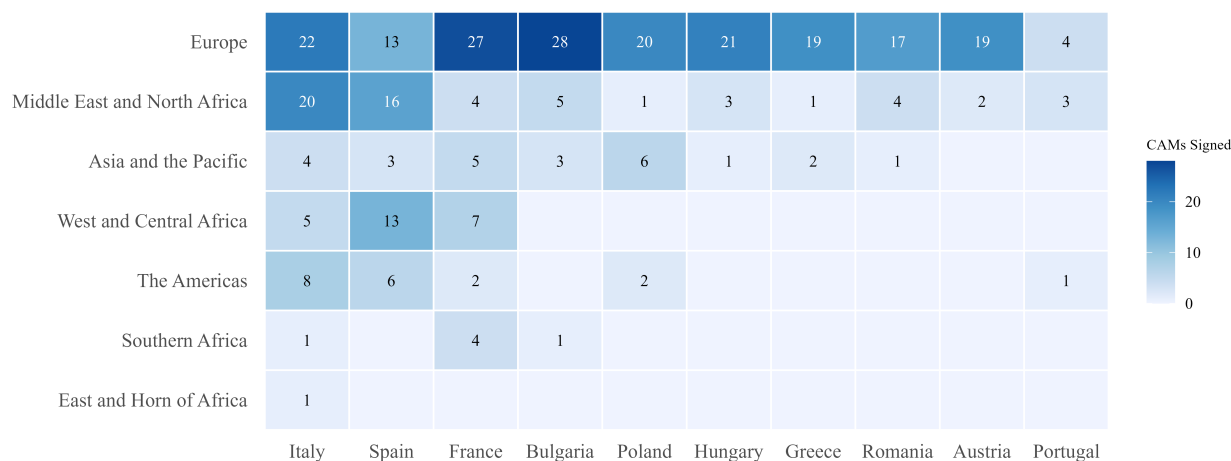


Figure 5: Number of CAMs signed by frontier state and UNHCR sub-region from 1990 to 2023. Empty cells indicate no CAMs signed between that frontier state and region.

capturing the credibility of migration as a political ‘threat’ independent of government type. I additionally control for migration-related protests as a measure of bottom-up political salience (Clark and Regan 2016), and the cumulative count of CAMs up to the observation year to account for path dependence in bilateral agreement formation.

Furthermore, domestic migration policy may independently predict foreign policy co-optation, incorporating deeper domestic relationships and perceptions about migration. That is, frontier states with more restrictive domestic postures toward migrants are likely to pursue more restrictive bilateral arrangements as well. I control for this alignment between domestic and foreign policy using the level of access to social assistance for recognized refugees, ranging from no access to benefits equivalent to those of residents, as a proxy for the underlying migration policy preferences of the frontier state Römer et al. (2021). Scholars have documented that EU membership shapes both the form and frequency of migration cooperation at different levels, operating through channels distinct from domestic electoral incentives. As such, I also control for EU membership, as the shared regulatory framework on migration among member states may independently influence the propensity to enact CAMs.³⁴ I control for cabinet ideological cohesion (*Ideologically Connected Cabinet*), as in-

34. Variable summaries are available in Appendix A.1.

ternally aligned cabinets may more readily pursue foreign policy instruments independent of electoral pressure as ideologically different parties find it difficult to land on a policy (Walther, Hellström, and Bergman 2019; Müller and Strøm 1999; Hellström et al. 2024b, 2024a). A control for cabinet incumbency as long-serving governments may have had greater opportunity to enact CAMs independent of electoral incentives, as their long tenure could imply (Hellström et al. 2024b, 2024a).

Lastly, I include dyadic controls to account for structural characteristics of the bilateral relationship that could influence the formation of CAMs. Following standard gravity model conventions, I control for contiguous borders, prior colonial relationships, and logged bilateral distance (*Contiguous Borders*, *Previous Colony*, *Distance (Log)*) (Mayer and Zignago 2011). This approach could be replicated with dyad fixed effects, though doing so would absorb the time-invariant dyadic variables of theoretical interest, precluding estimation of their individual effects. I include logged GDP per capita (*GDPPC (Log)*) as a proxy for the partner state’s economic development to account for state capacity and the economic drivers of emigration. I include partner’s life expectancy (in years) to proxy for humanitarian conditions, testing whether frontier states are more likely to enact CAMs with partners facing acute development challenges translating into lower life expectancy (The World Bank 2024). Similarly interested in the domestic context of the partner state, the partner state’s freedom of foreign movement and electoral democracy index (*Foreign Movement Index*, *Partner Non-Democratic*) are included to measure whether the partner’s treatment of citizens, migrants, and emigrants independently influences the frontier state’s propensity to enact CAMs (Coppedge et al. 2021).

Results

If foreign policy co-optation operates as theorized, mainstream governments should respond to rising right-wing populist seat share by enacting more CAMs, with the effect most

pronounced under left-of-center governments facing the greatest narrative pressure. Table 1 presents the effects of the interaction between the binary *Mainstream Government* indicator and continuous *RWP Seat Share* to test this prediction, with subsequent tables disaggregating by government type to test my second conjecture.

Model 1 includes only *Refugee Stock* and *Asylum Applications (Log)* to establish that objective migration pressure alone does not robustly predict CAM formation consistent with the theoretical expectation that co-optation is driven by electoral incentives rather than policy need. As the theory anticipates, migration crises are partly politically constructed. Governments and challengers can amplify the salience of any given migration flow to serve electoral ends, meaning that raw refugee or asylum figures are poor proxies for the political pressure that drives co-optation (see Figure 3). Model 2 includes the main explanatory variables alongside refugee stock and asylum applications, while models 3–5 incorporate full controls and fixed effects. Models 4 and 5 forward the dependent variable by one and two years respectively to assess whether the effect of right-wing populist pressure on CAM formation persists beyond the contemporaneous observation year. Coefficients from PPML models can be interpreted similarly to log-linear regressions. In models with categorical variables (*PM*, *PM + Cabinet*) and a continuous variable (*RWP Seat Share*), the interpretation of coefficients follows a semi-elasticity logic, with the main effect of right-wing populist seat share capturing its impact on the expected count of CAMs.³⁵ A coarse binary indicator of mainstream governance is also presented in Appendix C.2, alongside a collapsed left-right measure as additional robustness checks. The appendix discusses interpretation of the lagged specification and known estimation challenges with binary treatments in PPML models with country fixed effects.

Results show support for my first hypothesis across models 3–5, with these reaching

35. For log-transformed variables (*Refugee Stock (Log)*), coefficients approximate elasticities as in most gravity models. This means that a 1 percent increase changes the expected CAM count by approximately the outcome in percent. For binary variables (*EU Member*), coefficients can be interpreted by the exponentiation the coefficient ($e^{\beta} - 1 \times 100$) when switching from 0 to 1, whereas count coefficients (*Migration Protests*) follow same equation but percent changes by each one-unit increase.

statistical significance. Additional results from Table 2 also show support for my third hypothesis relating to the non-democratic status of the partner. There seems little evidence of the democratic status of partner country on the propensity for CAMs. Table 1 and Table 2 are in a similar structure but I ignore the statistically insignificant controls for visibility. Results for Table 1 use the main independent variable *PM*, which is government type constructed solely from the PM’s family, show support for my second hypothesis as well as displaying additional support for my first and third hypotheses. The interaction between RWP Seat Share and Socialist governments is positive and statistically significant across models 3 through 5, with the largest effect observed in the one-year forward lag. The interaction between RWP Seat Share and Social Democratic governments is similarly positive and significant across models 3 through 5, though the magnitude is considerably smaller than under Socialist governments. These results are again consistent with the prediction that ideological distance from right-wing challengers amplifies the co-optation effect.

Table 2 replicates this pattern using the PM + Cabinet measure, which incorporates the full partisan composition of the governing coalition, and yields substantively identical results. This pattern holds consistently across specifications with forward-lagged dependent variables ($t+1$, $t+2$). Overall, the most pronounced co-optation effects occur under left-wing governments, which enact significantly more CAMs as right-wing populist seat share rises, a pattern that holds across government type measures.

Table 1: Full tables available in Appendix B.1. FE PPML models measuring the count of total CAMs, given the party family of the PM interacted with right-wing populist seat share. *Center Right* government used as reference category.

	<i>Dependent Variable</i>				
	(1) CAMs (<i>mig</i>)	(2) CAMs (<i>basic</i>)	(3) CAMs (<i>t</i>)	(4) CAMs (<i>t</i> + 1)	(5) CAMs (<i>t</i> + 2)
<u><i>Main Explanatory Variables</i></u>					
Conservative		-0.52 (0.34)	-0.54 (0.56)	-0.78 (0.46)	-0.60 (0.49)
Soc Democratic		0.08 (0.29)	-0.21 (0.29)	-0.35 (0.34)	-0.007 (0.44)
Socialist		0.03 (0.40)	-9.8*** (2.5)	-89.2*** (3.4)	-10.4*** (2.1)
RWP Seat Share		-0.01 (0.008)	-0.06* (0.03)	-0.05 (0.03)	-0.07* (0.03)
× <i>Conservative</i>		0.02* (0.008)	0.05* (0.02)	0.06** (0.02)	0.06* (0.03)
× <i>Soc Democratic</i>		-0.004 (0.01)	0.03* (0.01)	0.04* (0.02)	0.03 (0.02)
× <i>Socialist</i>		0.009 (0.03)	0.49*** (0.11)	3.7*** (0.15)	0.52*** (0.08)
<u><i>Selected Control Variables</i></u>					
Refugee Stock (Log)	-0.04 (0.16)	-0.01 (0.14)	0.11 (0.18)	0.05 (0.17)	0.16 (0.22)
Asylum Applications (Log)	0.07 (0.05)	0.10* (0.05)	0.12 (0.07)	-0.005 (0.10)	0.12 (0.10)
Partner Non-Democratic			-0.31* (0.13)	-0.31** (0.10)	-0.22* (0.09)
Distance (Log)			-0.29** (0.10)	-0.23*** (0.07)	-0.06 (0.07)
Migration Protests			-0.07 (0.04)	-0.13* (0.06)	-0.17** (0.06)
Ideologically Connected Cabinet			-0.48*** (0.10)	-0.52*** (0.13)	-0.51** (0.18)
Observations	7,790	7,745	5,292	5,341	5,341
Squared Correlation	0.01137	0.01459	0.07350	0.03025	0.02100
Pseudo R ²	0.03607	0.04276	0.09742	0.06467	0.06000
BIC	2,564.9	2,588.6	1,812.1	1,849.3	1,832.1
Country fixed effects	✓	✓	✓	✓	✓
Decade fixed effects	✓	✓	✓	✓	✓
Controls included			✓	✓	✓

Notes: Clustered (country) standard errors in parentheses. Signif. codes: ***: 0.001, **: 0.01, *: 0.05.

Table 2: Full tables available in Appendix B.2. FE PPML models measuring the count of total CAMs, given the party composition of the government interacted with right-wing populist seat share. *Center Right* government used as reference category.

	<i>Dependent Variable</i>				
	(1) CAMs (<i>mig</i>)	(2) CAMs (<i>basic</i>)	(3) CAMs (<i>t</i>)	(4) CAMs (<i>t</i> + 1)	(5) CAMs (<i>t</i> + 2)
<u><i>Main Explanatory Variables</i></u>					
Conservative		-0.62 (0.40)	-0.59 (0.60)	-0.86 (0.55)	-0.46 (0.56)
Soc Democratic		-0.06 (0.23)	-0.27 (0.29)	-0.45 (0.38)	0.07 (0.41)
Socialist		-0.02 (0.40)	-9.7*** (2.3)	-88.7*** (3.1)	-9.6*** (1.7)
RWP Seat Share		-0.01 (0.007)	-0.05* (0.02)	-0.04 (0.02)	-0.06* (0.03)
× <i>Conservative</i>		0.02* (0.009)	0.05* (0.02)	0.05** (0.02)	0.05* (0.02)
× <i>Soc Democratic</i>		-0.005 (0.008)	0.03* (0.01)	0.04* (0.02)	0.02 (0.02)
× <i>Socialist</i>		0.0008 (0.03)	0.48*** (0.11)	3.7*** (0.13)	0.48*** (0.07)
<u><i>Selected Control Variables</i></u>					
Refugee Stock (Log)	-0.04 (0.16)	-0.0005 (0.14)	0.11 (0.19)	0.03 (0.18)	0.14 (0.21)
Asylum Applications (Log)	0.07 (0.05)	0.10* (0.04)	0.13 (0.06)	-0.004 (0.10)	0.12 (0.09)
Partner Non-Democratic			-0.30* (0.13)	-0.28** (0.09)	-0.15 (0.13)
Distance (Log)			-0.29** (0.10)	-0.22** (0.07)	-0.06 (0.06)
Migration Protests			-0.07* (0.04)	-0.16* (0.07)	-0.22*** (0.06)
Ideologically Connected Cabinet			-0.48*** (0.10)	-0.52*** (0.12)	-0.51*** (0.16)
Observations	7,811	7,811	5,329	5,378	5,378
Squared Correlation	0.01143	0.01468	0.07154	0.02874	0.02063
Pseudo R ²	0.03633	0.04280	0.09454	0.05679	0.05725
BIC	2,565.8	2,621.2	1,834.3	1,898.8	1,866.8
Country fixed effects	✓	✓	✓	✓	✓
Decade fixed effects	✓	✓	✓	✓	✓
Controls included			✓	✓	✓

Notes: Clustered (country) standard errors in parentheses. Signif. codes: ***: 0.001, **: 0.01, *: 0.05.

Overall, results indicate findings are in line with theoretical expectations. First, the

interaction between right-wing seat share is positively and significantly associated to the enactment of CAMs when interacted with mainstream governments. Even when taking into consideration asylum applications, refugee stocks, and election effects, these effects alone do not robustly predict more CAMs once other factors are in play. Second, the further left a government is the more correlated they will be with CAMs, indicating support for my second hypothesis as these governments face the stronger brunt from the narrative. The results further indicate that governments across the board enact CAMs, but they hold preferences over the partners they engage with. *Partner Non-Democratic* is negative and statistically significant across models 3 through 5 in Table 1, and in models 3 and 4 in Table 2. All else equal, including partner development and freedom of movement, a non-democratic partner is associated with a roughly 20 to 27 percent lower expected count of CAMs per dyad-year than a democratic partner (Appendix B.1, B.2). This is consistent with the reputational costs that the literature on cooperation with illiberal partners anticipates (Hafner-Burton 2008). These costs nonetheless act as a discount rather than a deterrent as more than two thirds of CAMs are still enacted with non-democratic partners, and more than a third with closed autocracies. Partner regime type shapes how often frontier states cooperate with a given partner, but it does not keep mainstream governments from co-opting CAMs with non-democracies, in line with my third hypothesis.

Historical Right-Wing Challenger Pressure

How governments handle right-wing populist seat share, and the narrative pressure alongside it, can be further refined by looking at the level of pressure throughout the data. Countries will vary their ability to withstand right-wing populist parties' pressure in their parliament given the historical and political system. For example, Figure 7 displays CAMs and the share of right-wing populist seat share across each cabinet (after each election) with some having a usually higher threshold for right-wing populist parties such as France and Spain, compared to others like Romania and Greece. Yet, although most countries display

an enactment of CAMs when countries are challenged by right-wing populist parties but these parties haven't taken control of government, their endogenous resistance to right-wing populist pressure is necessary to paint a clearer, more appropriate picture of mainstream governments. To identify whether CAM formation under mainstream governments is conditional on the intensity such pressure, I estimate PPML models interacting *PM + Cabinet* with indicator variables for right-wing populist seat share above or below the mean, up to one standard deviation (above or below) in the distribution.

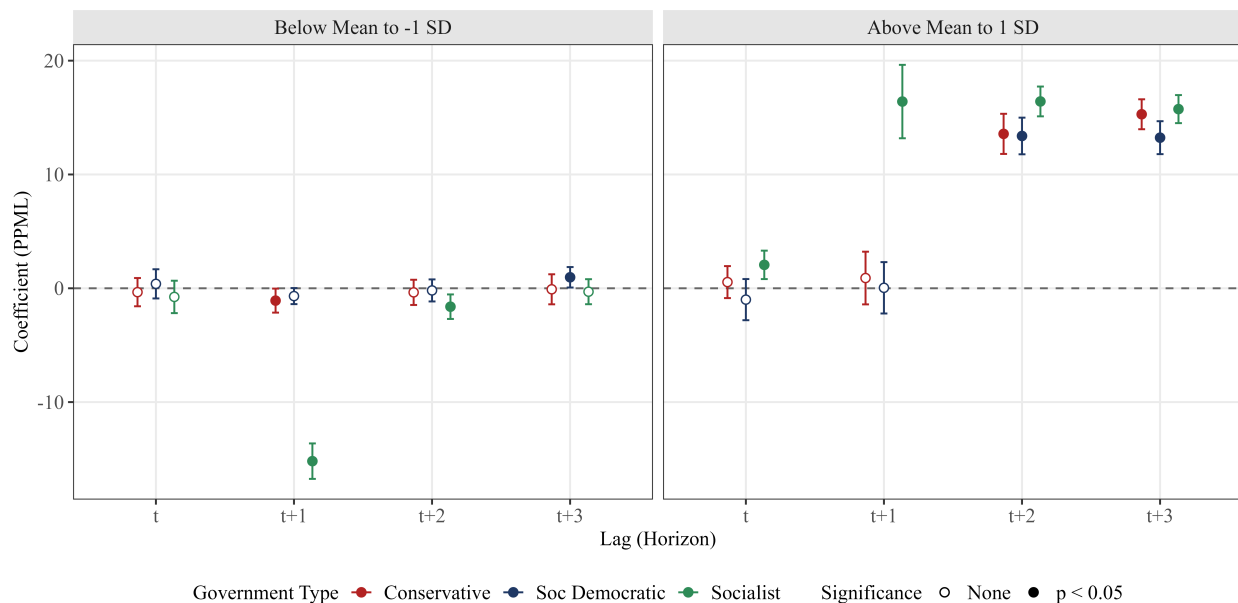


Figure 6: Coefficient estimates for the RWP seat-share bin interactions, by government type across the contemporaneous period and three annual leads. Filled points indicate significance at $p < 0.05$. Tables and additional discussion in Appendix B.3

These thresholds (mean, standard deviations) are calculated dynamically up to the corresponding cabinet to reflect the government's evaluation of right-wing populist seat share up to that point. Mainstream governments respond to right-wing threats only when these become electorally salient even across a wide range of variation. For instance, Italy's second and third Prodi governments were between Berlusconi-led governments. While the final Berlusconi government working closely with right-wing populist parties would later face backlash for its controversial migration deals with illiberal regimes, the Prodi cabinets, governing

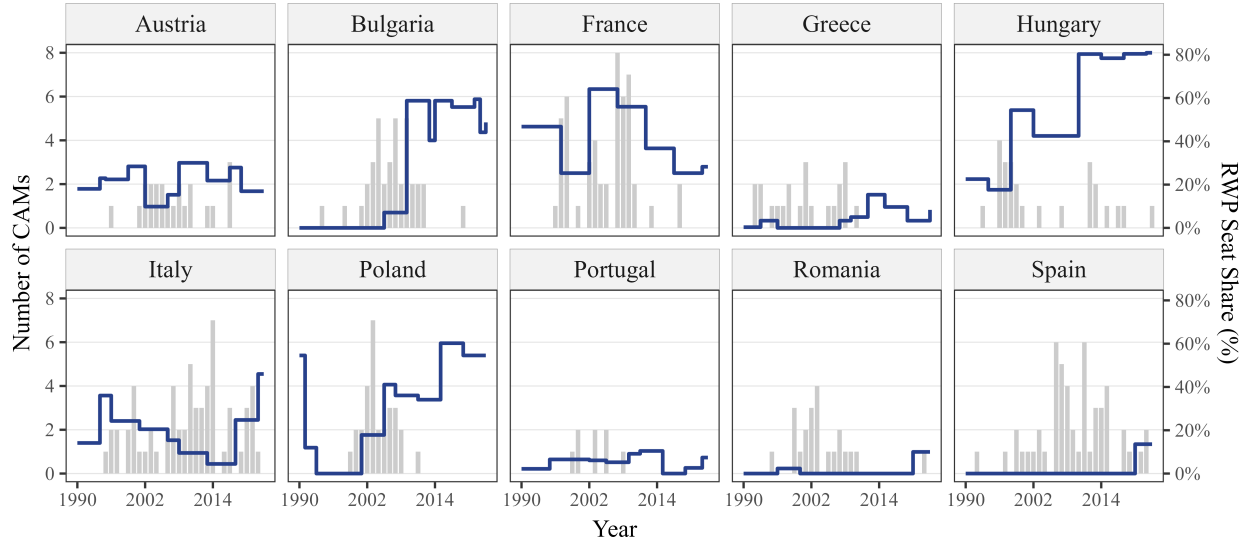


Figure 7: Outgoing CAMs are displayed as gray bars (left axis) alongside each country’s right-wing populist seat share as blue lines (right axis), from 1990 to 2023. The left axis is capped at 8 CAMs per country-year for legibility across panels.

earlier, could only assess the precedent set by Berlusconi’s first term. Results in Figure B.11 show that CAM formation significantly increases only when right-wing populist pressure is above one standard deviation from the mean, particularly under left-wing governments. I extend the time to $t + 3$ to observe if these patterns stay over a longer period of time or decrease. There is no significant positive effect when right-wing populist pressure is moderate or below average. In fact, quite the opposite is found in Figure B.11’s third model. This finding aligns with the logic of strategic co-optation in general. Additional patterns in Greece and Romania display that the enactment of CAMs could have warded off some increase of right-wing populist seat share, but such interpretation should still be taken lightly.³⁶

Conclusion

The study set out to explain international migration cooperation and its determinants across Western democracies utilizing a novel dataset. The study first presents a theory of foreign policy co-optation to clarify which type of foreign policy mainstream governments

36. Outcome placebo test using bilateral labor agreements or BLAs can be found in Appendix C.

pursue amid ongoing and rising right-wing populist threat. This is achieved by dissecting the type of government in power and by accounting for several predictors of migration cooperation—of which simple migration pressure was not enough to predict CAMs. The study additionally suggests how Western democratic states are not dissuaded by less pluralistic regimes when creating bilateral migration arrangements. I find that as right-wing populist parties gain electoral ground, mainstream governments respond by increasing the enactment of CAMs. These effects are most pronounced for left-of-center and moderate governments, which face the greatest need to reclaim narrative control over migration from right-wing challengers. Last, contrary to common expectations that regime type would moderate partnerships, the findings suggest that partner regime types have little influence in the domestic electoral calculus of mainstream governments.

The findings carry broader implications for research on international order, cooperation, and migration. The dataset provides a foundation for future research on border externalization and migrant safety by systematically cataloging CAMs. Future research can extend this analysis to other contexts, including the United States, Commonwealth democracies, and emerging democracies, where similar dynamics may operate under different conditions. However, theorization must move beyond CAMs to capture the broader militarization of humanitarian migration cooperation that has developed over the last thirty years, especially agreements led by Global South countries, such as tripartite repatriation arrangements and refugee-centered peace agreements.

Furthermore, the theory of foreign policy co-optation presented can extend to other policy domains. For example, in contexts where left-wing populist parties dominate the challenger space, as in parts of Latin America, mainstream governments might engage in analogous foreign policy co-optation through bilateral development or labor agreements that appropriate the challenger’s narrative. Even in two-party systems, where parties are often concerned about differentiating themselves and large programmatic shifts, we are likely to still observe co-optation of narrative-important policies when political survival is at stake.

The study showed how migration policies aligned with right-wing populist standards can extend into the international arena, reshaping patterns of cooperation in ways that may undermine or reshape global democratic norms, and create interdependence and reliance of extractive actors. This naturally raises a concern over the future of migration management beyond bilateral agreements. If democratic governments ignore the types of policies they develop to stay in power, their engagement is likely to cause migrant suffering and foment detachment. Yet, if these government choose to ignore their domestic challengers, this poses a threat to the stability of mainstream politics. This study acknowledges the difficulty of this dilemma and offers an observation.

Bilateral and security policy on migration will naturally rely on less-pluralist partners to achieve objectives. However, this reliance does not require the absence of protection for migrants and vulnerable individuals. Although this study does not analyze conditionalities within CAMs, preliminary evidence shows that humanitarian clauses, references to international law or good governance, protections for women and children, and rescue provisions are largely absent.³⁷ Amending these and future agreements offers a kernel of hope to counter the erosion of liberal norms that Western governments disavow in practice. Executive discretion in diplomatic negotiations especially creates space to embed these provisions, but it will require clarity and credible threat of punishment for mishandling enforcement.

The international liberal order facing deep contestation signals the need to identify and modify practices that entrench illiberal international norms. The findings show that right-wing populist pressure enables executives to manage electoral opposition through foreign policy. Practices that can shape protections for the more than 123 million forcibly displaced people worldwide, the highest ever reached in modern history. This highlights the need to revisit prevailing assumptions about how domestic politics shape international migration cooperation and the international order.

37. Subsequent work examines this conditionality across partner regions and its evolution.

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Appendix

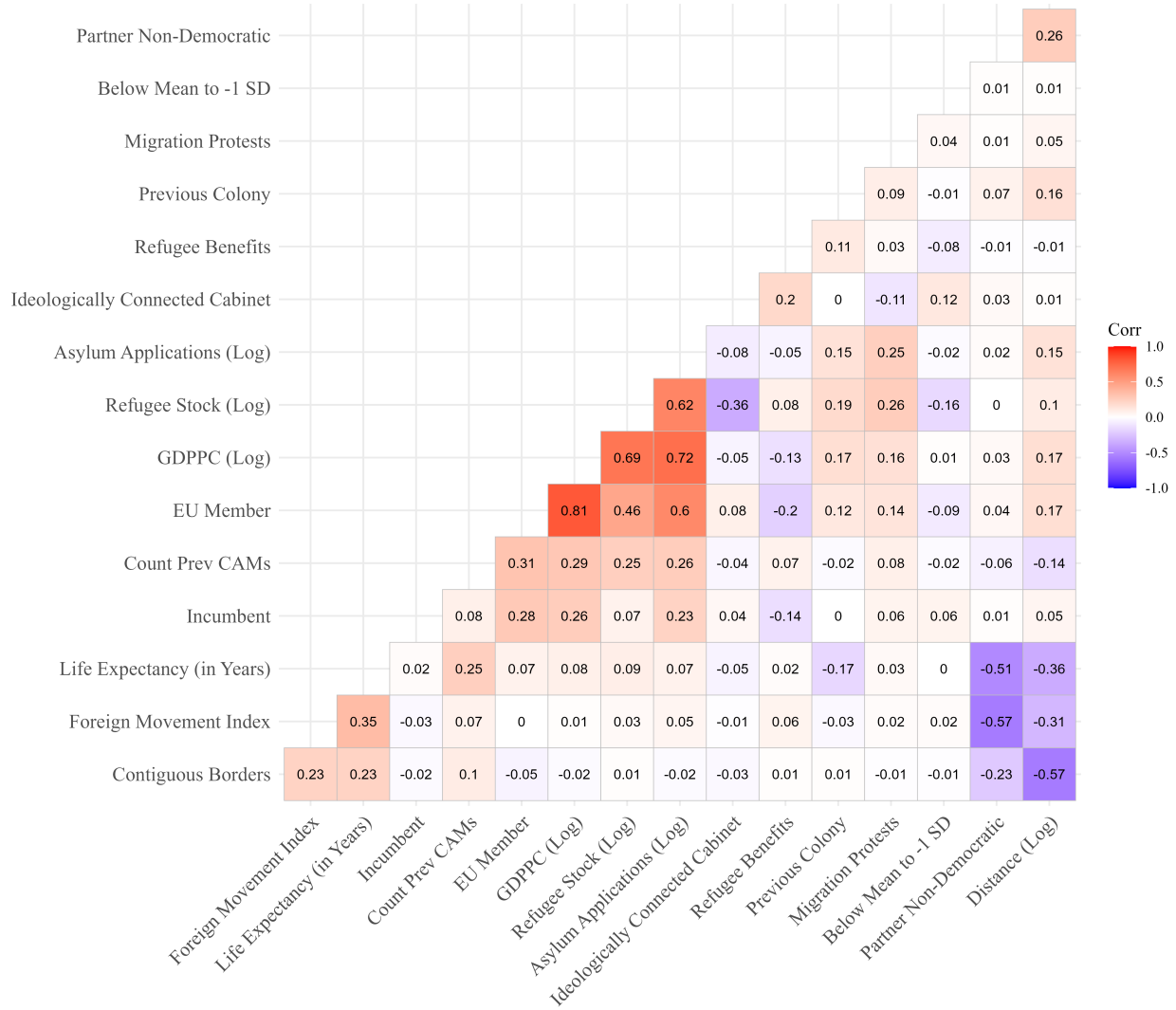
A	General Supplements and Additional Models	2
A.1	Summary Statistics	2
A.2	Correlation Plot	3
A.3	Mean Seats Plot and Cutoffs	4
A.4	Data Missingness Plot	6
A.5	CAMs Signing and Enforcement	7
A.6	Manual Party Additions	8
A.7	OLS Models	10
B	PPML Complete Tables Models	12
B.1	PPML PM Family	12
B.2	PPML PM and Cabinet Composition	14
B.3	PPML Government Type and Mean RWP Presence	15
C	Robustness Checks Using BLAs, Broader Categorization	19
C.1	Bilateral Labor Arrangements (BLAs) on Migration	19
C.2	Binary and Simple Left Measurement	21

A General Supplements and Additional Models

A.1 Summary Statistics

Variable	N	Mean	Std. Dev.	Min	Pctl. 25	Pctl. 75	Max
CAM Rows	9493	0.035	0.18	0	0	0	1
Mainstream Government	9493	0.97	0.17	0	1	1	1
RWP Seat Share	9493	25	20	0	6.1	41	64
Partner Non-Democratic	9165	0.67	0.47	0	0	1	1
Foreign Movement Index	9165	1.4	1.1	-3.6	0.9	2.3	2.9
Life Expectancy (in Years)	8031	73	6.6	42	70	77	86
Contiguous Borders	9255	0.15	0.35	0	0	0	1
Distance (Log)	9255	2743	2715	60	812	3763	13002
Previous Colony	9255	0.092	0.29	0	0	0	1
Refugee Stock (Log)	9386	9.7	1.9	5.3	8.4	12	14
Asylum Applications (Log)	8628	8.5	2.5	0	7.5	10	12
Migration Protests	9465	0.44	1.5	0	0	0	18
GDP Per Capita (Log)	9465	9.7	0.93	7	9.3	10	11
Ideologically Connected Cabinet	9465	0.46	0.5	0	0	1	1
Decade	9465						
... 1980s	350	4%					
... 1990s	2812	30%					
... 2000s	3000	32%					
... 2010s	2630	28%					
... 2020s	673	7%					
Refugee Benefits	7741	0.75	0.42	0	0.5	1	1
EU Member	8907	0.82	0.39	0	1	1	1
Count Prev CAMs	9493	0.59	0.68	0	0	1	11
Incumbent	9493	0.46	0.5	0	0	1	1

A.2 Correlation Plot



The plot presents pairwise correlations among the principal predictors included in the dyadic analyses. Overall, most coefficients are modest in magnitude, with the majority clustered near zero. Distance remains only weakly correlated with refugee stocks (0.10) and asylum applications (0.15), though it correlates more substantially with contiguous borders (-0.57), foreign-movement freedoms (-0.31), and partner life expectancy (-0.36). This seems to reflect the geographic clustering of these variables rather than a migration-specific relationship.

The strongest positive associations arise among structural migration and development variables. EU membership is strongly correlated with GDP per capita (0.81), and moderately correlated with larger refugee stocks (0.46) and greater numbers of asylum applications (0.60), reflecting typical destination-country characteristics. GDP per capita is similarly correlated with refugee stocks (0.69) and asylum applications (0.72), while refugee stocks and

asylum applications are themselves correlated with one another as expected (0.62). By contrast, GDP per capita shows essentially no association with foreign-movement freedoms (0.01), unlike what a purely economic-openness account might predict. Lastly, some moderate correlations emerge among regime and development indicators. In particular, partner non-democracy is negatively associated with life expectancy (-0.51) and foreign-movement freedoms (-0.57), consistent overlap given lower ability of movement and expression in non-democracies.

A.3 Mean Seats Plot and Cutoffs

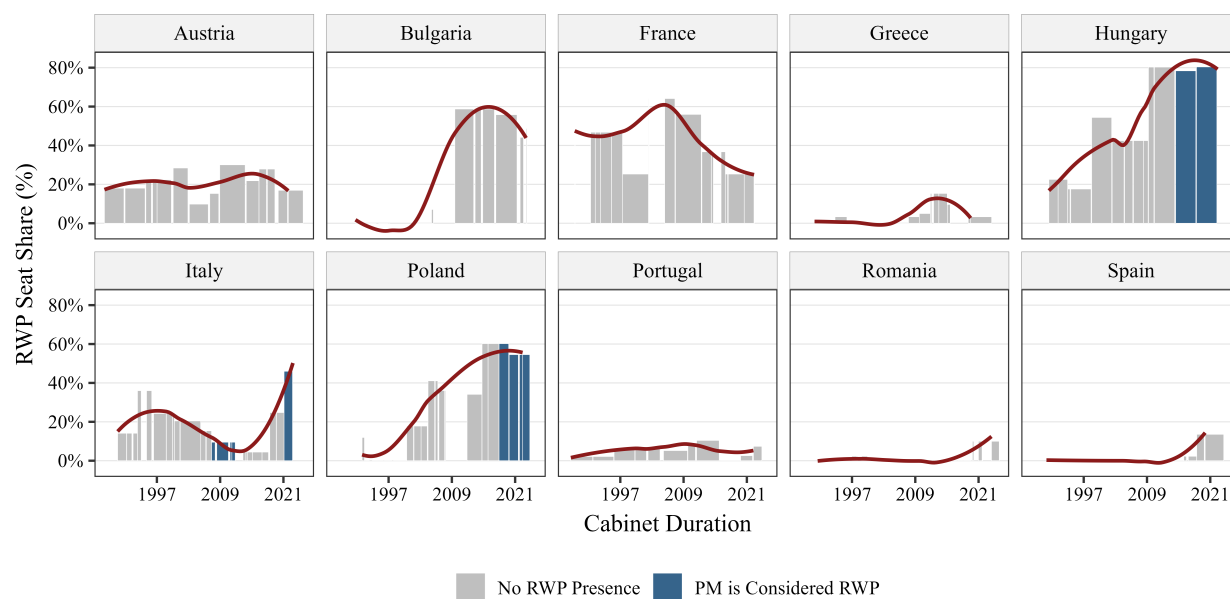


Figure A.8: Running average of RWP seat share by origin country. Each panel is a frontier's cumulative mean up through each cabinet, with the navy color indicating that the PM of that government is considered a right-wing populist leader.

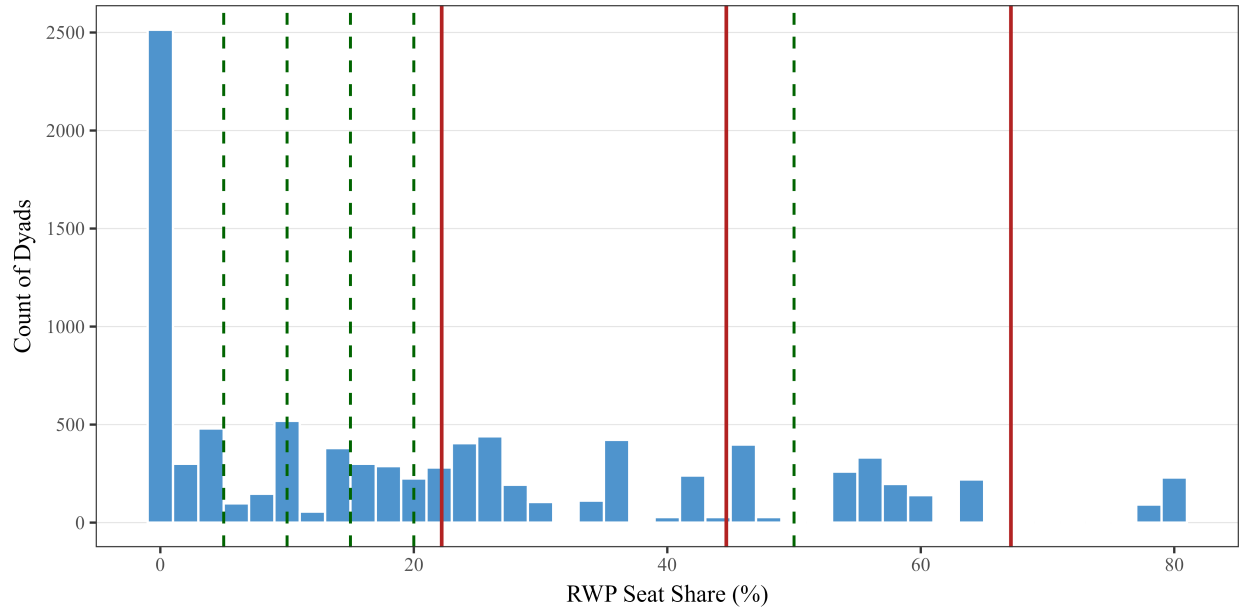
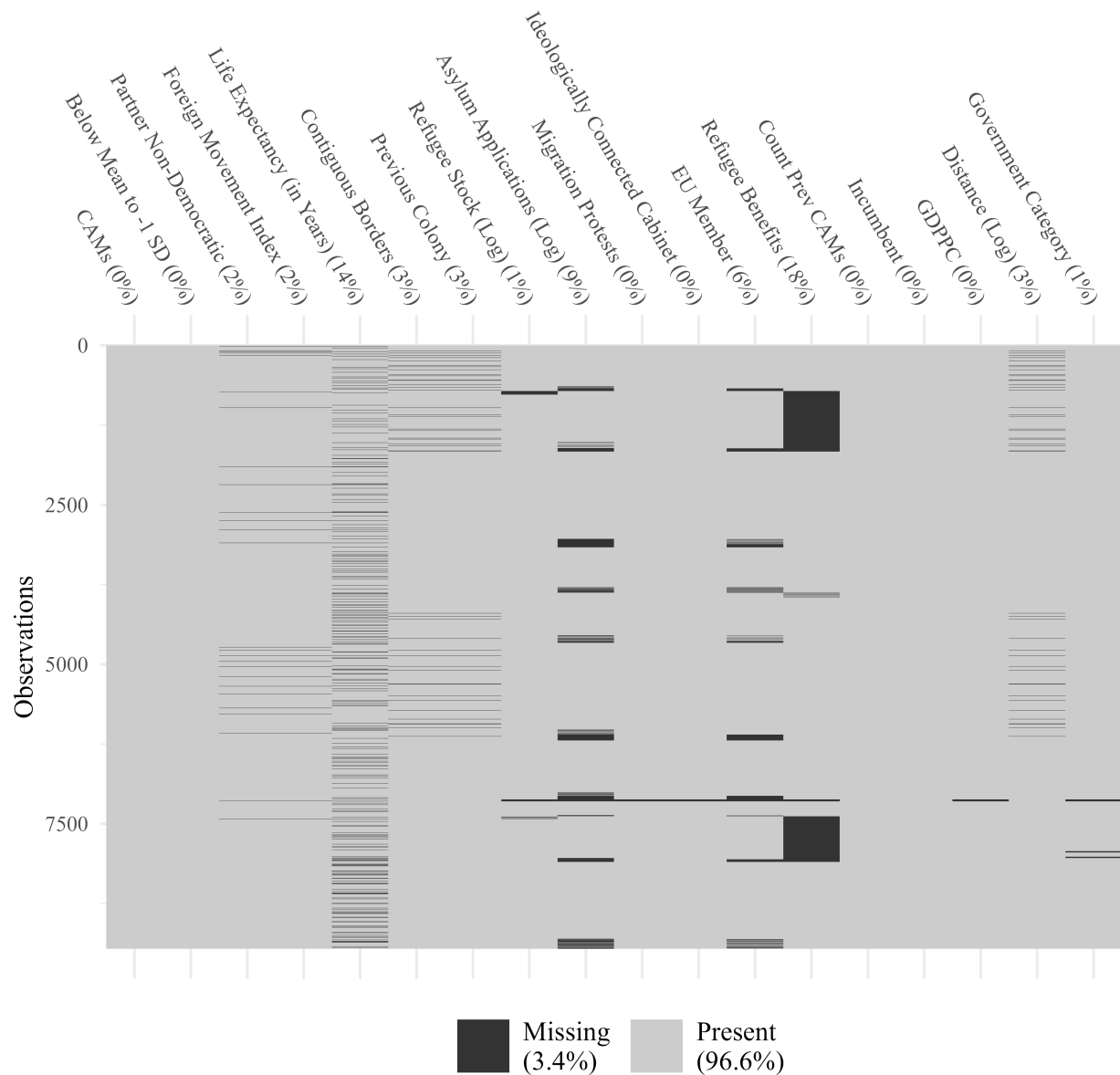


Figure A.9: Distribution of RWP seat share with thresholds. Green dashed line indicates fixed cutoffs; Red solid lines indicate Mean + 1 standard deviation, and Mean + 2 standard deviations.”

A.4 Data Missingness Plot



The plot showcases the missingness structure for all variables included in the dyadic dataset. Overall, missing data are limited to only 3.4 percent of all observations, and most variables exhibit near-complete coverage. The variables with the highest missingness are *Refugee Benefits*, *Life Expectancy*, and *Asylum Applications (Log)*, although in all cases the proportion of missing entries remains low. Missingness appears scattered rather than clustered in systematic blocks, suggesting that gaps are primarily due to country-year availability, and the overall low rate of absence minimizes concerns about bias arising from incomplete data.

A.5 CAMs Signing and Enforcement

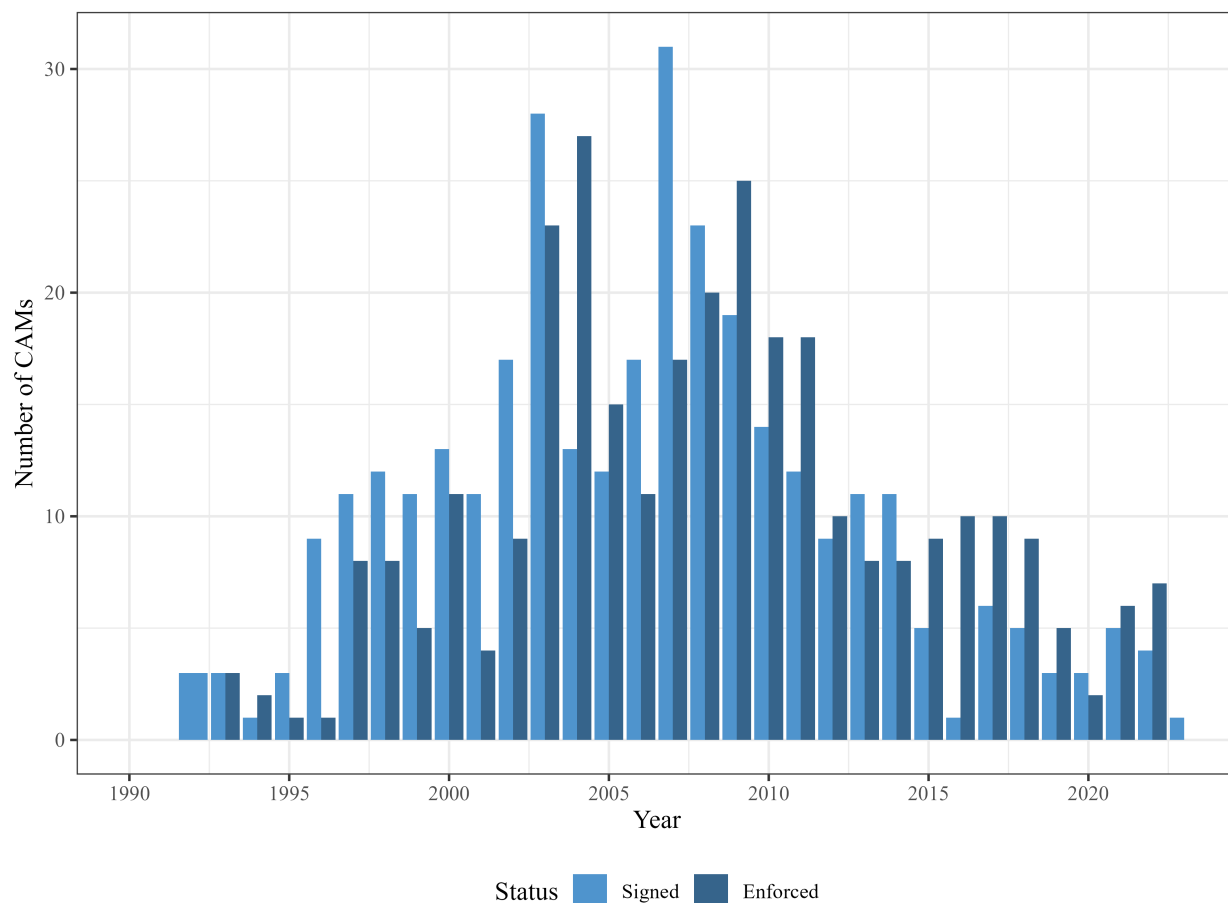


Figure A.10: Total CAMs signed and enforced by year across all frontier states with all partner countries. Note that there are years when CAMs are signed but enforcement begins at a different year.

CAMs date of signature and date of enforcement are displayed in the figure, with the dark bars showing the number of CAMs signed each year from 1990 to 2023, while the lighter bars show the number enforced in the same year. Many agreements are enforced later than the year in which they are signed, with only two CAMs having no recorded enforcement date. The pattern shows a rapid build-up of CAM activity from the mid-1990s, with pronounced waves of signings and enforcements in the late 1990s and mid-2000s, followed by a gradual decline after 2010. The summary statistics underline this timing gap: about 20 percent of CAMs are enforced within 6 months, about 42 percent of CAMs are enforced within one year of signature, roughly 63 percent within two years, 78 percent within three, and nearly 87 percent within four years, while the remaining agreements either take longer to enter into force or currently lack a documented enforcement date.

A.6 Manual Party Additions

To improve the accuracy of party-family classifications in the CEEWED dataset, I conducted post-hoc manual recoding for parties with missing or ambiguous family assignments. These corrections were based on secondary sources including the Manifesto Project Handbook, the Party Government and Election Database Codebook, and additional qualitative research for selected country cases. For each country, I identified parties with missing family values and assigned them to the appropriate ideological family (e.g., Conservative, Liberal, Green, Socialist) using their programmatic positions, electoral alliances, and expert evaluations which would then be refined through my rule-based categorization, found in Table A.3. Examples include classifying Italy’s *Fratelli d’Italia* (FdI) and *Lega Nord* (LN) as Radical Right (70), Democratic Party (Dem.) as social democratic (40), Hungary’s *Fidesz* and KDNP as Radical Right (70), Bulgaria’s *We Continue the Change* (PP) as Liberal (40), Croatia’s *Homeland Movement* (DP) as Radical Right (70), Romania’s *USR PLUS* as Liberal (40), and Spain’s *Vox* as Radical Right (70) while retaining *Partido Popular* (PP) as mainstream Conservative (60). The higher the number, the more right-leaning these are considered up to (80) where these happen to be within the ‘Special’ category of my coding. These manual corrections ensure fuller coverage and greater conceptual validity of the party-family measure, particularly for newer, merged, or region-specific parties not fully captured by standard datasets. Table A.4 documents a second, cabinet-year-level layer of manual correction. In cases where the party-level recoding in Table A.3 was not sufficient to fully propagate through the government classification function, I applied direct overrides to the PM-based and cabinet-based government categories for the affected country-years for ‘perceived’ right-wing prime ministers—attempting to avoid presentism as much as possible.

Table A.3: Manual party family coding by country, party family in parentheses.

Bulgaria	Croatia	France	Greece	Hungary	Italy	Poland	Romania	Spain
DB (40)	DP (70)	OG (10)	ND (60)	KDNP (70)	FdI (70)	NL (30)	PLR (40)	PP (60)
PP (40)				Fidesz (70)	Dem. (40)	PL2050 (60)	USR PLUS (40)	Vox (70)
ITN (60)				Somogyért (90)	FLI (60)		UDMR (90)	
OP (70)				MSzP (30)	FI-PdL (60)		PDL (60)	
ZD (10)					LN (70)			
					UDR-UDEUR (50)			
					PR (20)			
					PT-CN (60)			
					Rep. (40)			

Table A.4: Manual overrides to PM-based and cabinet-based government classification, by country and year range.

Country	Years	Assigned Category	Basis
Italy	1995–1996	Special	Technical/caretaker government; PM/cabinet family missing
Poland	Any Instance	Soc Democratic	Reclassified after further research
Bulgaria	1997–2000	Conservative	PM/cabinet family missing; reclassified after research
Bulgaria	2023–present	Special	Cabinet family missing; reclassified after research
Hungary	2014–2023	RWP	Fidesz-led government reclassified as RWP
Italy	2008–2012; 2022–present	RWP	FdI/Lega-led coalition governments reclassified as RWP
Poland	2018–2023	RWP	PiS-led government reclassified as RWP

A.7 OLS Models

Table A.5: Fixed-effects OLS models estimating the count of CAMs signed by frontier states in a given year. This serves as a robustness check on the dyadic PPML specifications, the unit of observation is the country-year by focusing exclusively on the frontier/receiving state and moving away from partner-specific variation.

	CAMs		CAMs ($t+1$)	CAMs ($t+2$)
	(1)	(2)	(3)	(4)
Conservative	-0.74*	-0.25	-0.13	-0.56
	(0.29)	(0.44)	(0.53)	(0.65)
Soc Democratic	-0.08	0.19	0.06	-0.38
	(0.17)	(0.32)	(0.30)	(0.56)
Socialist	0.03	-1.2	-1.1	-0.36
	(0.31)	(0.67)	(0.88)	(0.71)
RWP Seat Share	-0.02*	-0.03	-0.03*	-0.01
	(0.007)	(0.02)	(0.01)	(0.01)
Refugee Stock (Log)	0.03	0.06	-0.02	-0.006
	(0.10)	(0.20)	(0.18)	(0.12)
Asylum Applications (Log)	0.08	0.08	-0.02	-0.05
	(0.04)	(0.08)	(0.11)	(0.14)
Conservative $\times$ RWP Seat Share	0.03**	0.04*	0.04**	0.02*
	(0.007)	(0.01)	(0.009)	(0.009)
Soc Democratic $\times$ RWP Seat Share	-0.001	0.01	0.02**	0.02
	(0.007)	(0.007)	(0.006)	(0.01)
Socialist $\times$ RWP Seat Share	0.003	0.13*	0.10	0.05
	(0.02)	(0.05)	(0.06)	(0.03)
Migration Protests		-0.08*	-0.02	0.09
		(0.03)	(0.03)	(0.05)
GDPPC (Log)		0.74	0.69	0.54
		(0.96)	(1.0)	(0.91)
Ideologically Connected Cabinet		-0.48	-0.50**	0.10
		(0.22)	(0.14)	(0.19)
EU Member		-0.78	-0.53	-0.58
		(0.69)	(0.63)	(0.63)
Refugee Benefits		-0.08	0.43	0.002
		(0.49)	(0.64)	(0.55)
Count Prev CAMs		-1.2*	-1.6*	-1.4
		(0.41)	(0.62)	(0.58)
Observations	277	219	219	219
R ²	0.25020	0.28943	0.29994	0.29703
Within R ²	0.03713	0.12100	0.13706	0.10972
Country fixed effects	✓	✓	✓	✓
Decade fixed effects	✓	✓	✓	✓

Clustered (Country) standard-errors in parentheses

*Signif. Codes: ***: 0.001, **: 0.01, *: 0.05, .: 0.1*

B PPML Complete Tables Models

B.1 PPML PM Family

Table B.6: PPML models measuring the count of total CAMs, given the party family of the PM interacted with right-wing populist seat share.

		CAMs		CAMs (t+1)	CAMs (t+2)
	(1)	(2)	(3)	(4)	(5)
Refugee Stock (Log)	-0.04 (0.16)	-0.01 (0.14)	0.11 (0.18)	0.05 (0.17)	0.16 (0.22)
Asylum Applications (Log)	0.07 (0.05)	0.10* (0.05)	0.12 (0.07)	-0.005 (0.10)	0.12 (0.10)
Conservative		-0.52 (0.34)	-0.54 (0.56)	-0.78 (0.46)	-0.60 (0.49)
Soc Democratic		0.08 (0.29)	-0.21 (0.29)	-0.35 (0.34)	-0.007 (0.44)
Socialist		0.03 (0.40)	-9.8*** (2.5)	-89.2*** (3.4)	-10.4*** (2.1)
RWP Seat Share		-0.01 (0.008)	-0.06* (0.03)	-0.05 (0.03)	-0.07* (0.03)
Conservative × RWP Seat Share		0.02* (0.008)	0.05* (0.02)	0.06** (0.02)	0.06* (0.03)
Soc Democratic × RWP Seat Share		-0.004 (0.01)	0.03* (0.01)	0.04* (0.02)	0.03 (0.02)
Socialist × RWP Seat Share		0.009 (0.03)	0.49*** (0.11)	3.7*** (0.15)	0.52*** (0.08)
Partner Non-Democratic			-0.31* (0.13)	-0.31** (0.10)	-0.22* (0.09)
Foreign Movement Index			-0.08 (0.06)	-0.08 (0.06)	-0.05 (0.04)
Life Expectancy (in Years)			-0.01 (0.02)	-0.02* (0.009)	-0.02* (0.007)
Contiguous Borders			-0.17 (0.30)	-0.22 (0.34)	-0.29 (0.31)
Distance (Log)			-0.29** (0.10)	-0.23*** (0.07)	-0.06 (0.07)
Previous Colony			0.004 (0.40)	-0.10 (0.19)	-0.20 (0.11)
Migration Protests			-0.07 (0.04)	-0.13* (0.06)	-0.17** (0.06)
GDPPC (Log)			0.87 (0.81)	0.22 (0.69)	-0.35 (0.61)
Ideologically Connected Cabinet			-0.48*** (0.10)	-0.52*** (0.13)	-0.51** (0.18)
EU Member			-0.74 (0.71)	-0.34 (0.56)	-0.10 (0.59)
Refugee Benefits			0.02 (1.0)	-0.32 (1.1)	-0.42 (1.1)
Count Prev CAMs			-1.7* (0.83)	-1.0** (0.33)	-0.60*** (0.10)
Observations	7,790	7,745	5,292	5,341	5,341
Squared Correlation	0.01137	0.01459	0.07350	0.03025	0.02100
Pseudo R ²	0.03607	0.04276	0.09742	0.06467	0.06000
BIC	2,564.9	2,588.6	1,812.1	1,849.3	1,832.1
Country fixed effects	✓	✓	✓	✓	✓
Decade fixed effects	✓	✓	✓	✓	✓

Clustered (Country) standard-errors in parentheses

*Signif. Codes: ***: 0.001, **: 0.01, *:0.05, .: 0.1*

B.2 PPML PM and Cabinet Composition

Table B.7: PPML models measuring the count of total CAMs, given the party composition of the government and PM interacted with right-wing populist seat share.

	(1)	CAMs (2)	(3)	CAMs (t+1) (4)	CAMs (t+2) (5)
Refugee Stock (Log)	-0.04 (0.16)	-0.0005 (0.14)	0.11 (0.19)	0.03 (0.18)	0.14 (0.21)
Asylum Applications (Log)	0.07 (0.05)	0.10* (0.04)	0.13 (0.06)	-0.004 (0.10)	0.12 (0.09)
Conservative		-0.62 (0.40)	-0.59 (0.60)	-0.86 (0.55)	-0.46 (0.56)
Soc Democratic		-0.06 (0.23)	-0.27 (0.29)	-0.45 (0.38)	0.07 (0.41)
Socialist		-0.02 (0.40)	-9.7*** (2.3)	-88.7*** (3.1)	-9.6*** (1.7)
RWP Seat Share		-0.01 (0.007)	-0.05* (0.02)	-0.04 (0.02)	-0.06* (0.03)
Incumbent		-0.18 (0.19)	-0.03 (0.17)	-0.04 (0.19)	-0.15 (0.25)
Conservative × RWP Seat Share		0.02* (0.009)	0.05* (0.02)	0.05** (0.02)	0.05* (0.02)
Soc Democratic × RWP Seat Share		-0.005 (0.008)	0.03* (0.01)	0.04* (0.02)	0.02 (0.02)
Socialist × RWP Seat Share		0.0008 (0.03)	0.48*** (0.11)	3.7*** (0.13)	0.48*** (0.07)
Partner Non-Democratic			-0.30* (0.13)	-0.28** (0.09)	-0.15 (0.13)
Foreign Movement Index			-0.08 (0.06)	-0.08 (0.06)	-0.04 (0.05)
Life Expectancy (in Years)			-0.01 (0.02)	-0.02 (0.01)	-0.01 (0.008)
Contiguous Borders			-0.17 (0.29)	-0.17 (0.30)	-0.31 (0.31)
Distance (Log)			-0.29** (0.10)	-0.22** (0.07)	-0.06 (0.06)
Previous Colony			0.01 (0.41)	-0.03 (0.25)	-0.19 (0.11)
Migration Protests			-0.07* (0.04)	-0.16* (0.07)	-0.22*** (0.06)
GDPPC (Log)			0.87 (0.82)	0.14 (0.72)	-0.33 (0.57)
Ideologically Connected Cabinet			-0.48*** (0.10)	-0.52*** (0.12)	-0.51*** (0.16)
EU Member			-0.74 (0.69)	-0.34 (0.54)	-0.11 (0.57)
Refugee Benefits			0.09 (0.95)	0.55 (0.80)	0.31 (0.91)
Count Prev CAMs			-1.6 (0.85)	-0.81 (0.44)	-0.58*** (0.10)
Observations	7,811	7,811	5,329	5,378	5,378
Squared Correlation	0.01143	0.01468	0.07154	0.02874	0.02063
Pseudo R ²	0.03633	0.04280	0.09454	0.05679	0.05725
BIC	2,565.8	2,621.2	1,834.3	1,898.8	1,866.8
Country fixed effects	✓	✓	✓	✓	✓
Decade fixed effects	✓	✓	✓	✓	✓

Clustered (Country) standard-errors in parentheses

*Signif. Codes: ***: 0.001, **: 0.01, *: 0.05, .: 0.1*

B.3 PPML Government Type and Mean RWP Presence

Table B.8 reports the full specification underlying the binned robustness check presented in the main text (Table ??). Rather than modeling right-wing populist (RWP) seat share as a continuous predictor, this specification splits the RWP distribution into two indicator variables relative to the country-specific running mean of *Below Mean to -1 SD* and *Above Mean to 1 SD*, each interacted with government type (Conservative, Soc Democratic, and Socialist, relative to Center Right). Models are estimated separately for each bin at the contemporaneous period and at three annual leads ($t+1$ through $t+3$), following the same fixed-effects PPML specification and control set as the main analysis. Figure B.11 plots the full set of interaction coefficients from Table B.8 directly, by government type, horizon, and bin.

The *Below Mean to -1 SD* results (columns 1, 3, 5, and 7, left panel of Figure B.11) are broadly consistent with the continuous specification reported in the main text. The coefficients are modest in magnitude and standard errors are well-behaved across all three government types and all four horizons, with the exception of the Socialist interaction at $t+1$, discussed below.

The *Above Mean to 1 SD* results (columns 2, 4, 6, and 8; right panel of Figure B.11) reflect a comparatively small subsample at longer horizons, few country-decade cells contain governments of a given type operating under RWP seat share a full standard deviation above the country mean. Beginning at $t+2$, the interaction terms for Conservative, Soc Democratic, and Socialist governments converge to a similar magnitude (13 to 16), each estimated with a tight standard error. This is the expected behavior of the PPML estimator when a cell's outcome exhibits little within-cell variation (Correia, Guimarães, and Zylkin 2019). Simply, the coefficient is not an artifact of misspecification but a direct consequence of few observations populating the *Above Mean to 1 SD* bin at these horizons, once further divided by government type. The pattern holds whether government type is coded from the prime minister's party alone or the combined (PM+Cabinet) measure, and whether Socialist is retained separately or merged with Soc Democratic, as found in Figure B.12. Figure B.13 places the *Above Mean to 1 SD* results from both specifications side by side with convergence toward the same magnitude at $t+2$ and $t+3$ persisting. This is whether Socialist is estimated separately or merged into a broader Left category, confirming that it is a property of the bin's sample size at longer leads, not of any particular variable construction.

Given this, the continuous specification reported in the main text, which does not require partitioning the sample by RWP threshold, is the estimate I report as the primary result for characterizing effects at high RWP seat share. The *Above Mean to 1 SD* coefficients at $t+2$ and $t+3$ are included here as part of the full robustness table but reflect the coefficient identified with the best available data at the time.

Table B.8: Fixed-effects PPML estimates of the count of CAMs at varying times, by government type interacted with right-wing populist seat share bins.

	CAMs		CAMs (t+1)		CAMs (t+2)		CAMs (t+3)	
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
Conservative	0.12 (0.38)	-0.15 (0.38)	0.41 (0.41)	0.05 (0.41)	0.06 (0.37)	-0.009 (0.45)	0.16 (0.47)	-0.30 (0.48)
Soc Democratic	-0.12 (0.41)	0.14 (0.31)	0.32 (0.37)	0.19 (0.34)	0.18 (0.39)	0.16 (0.37)	-0.17 (0.50)	0.14 (0.43)
Socialist	0.22 (0.64)	-1.4* (0.69)	0.79 (0.67)	-14.8*** (0.33)	0.61 (0.70)	-1.4 (0.89)	-0.37 (0.75)	-1.5* (0.71)
Below Mean to -1 SD	-0.51 (0.39)		0.91*** (0.18)		0.40 (0.21)		-0.65 (0.37)	
Partner Non-Democratic	-0.31* (0.15)	-0.33* (0.15)	-0.30** (0.10)	-0.31** (0.10)	-0.20* (0.10)	-0.20* (0.10)	-0.03 (0.18)	-0.04 (0.19)
Foreign Movement Index	-0.08 (0.05)	-0.08 (0.05)	-0.08 (0.06)	-0.08 (0.06)	-0.05 (0.04)	-0.05 (0.04)	-0.09 (0.06)	-0.10 (0.06)
Life Expectancy (in Years)	-0.01 (0.02)	-0.01 (0.02)	-0.02* (0.009)	-0.02* (0.008)	-0.01 (0.008)	-0.01* (0.008)	-0.01 (0.01)	-0.01 (0.01)
Contiguous Borders	-0.16 (0.30)	-0.15 (0.29)	-0.21 (0.32)	-0.20 (0.33)	-0.29 (0.29)	-0.28 (0.30)	-0.02 (0.16)	-0.02 (0.16)
Distance (Log)	-0.28** (0.10)	-0.29** (0.10)	-0.23*** (0.07)	-0.23*** (0.07)	-0.04 (0.06)	-0.05 (0.06)	-0.09* (0.04)	-0.10* (0.04)
Previous Colony	-0.003 (0.40)	-0.008 (0.41)	-0.10 (0.18)	-0.10 (0.18)	-0.19 (0.11)	-0.19 (0.11)	-0.11 (0.19)	-0.11 (0.19)
Refugee Stock (Log)	-0.08 (0.19)	0.007 (0.18)	-0.03 (0.15)	-0.007 (0.15)	0.06 (0.22)	0.10 (0.22)	-0.25 (0.17)	-0.16 (0.19)
Asylum Applications (Log)	0.11 (0.07)	0.16* (0.07)	-0.005 (0.08)	0.02 (0.09)	0.11 (0.08)	0.10 (0.08)	0.04 (0.09)	0.03 (0.09)
Migration Protests	-0.09 (0.06)	-0.08 (0.04)	-0.14** (0.06)	-0.13* (0.06)	-0.20*** (0.04)	-0.19** (0.06)	-0.08 (0.04)	-0.06* (0.03)
GDPPC (Log)	1.2 (0.75)	0.84 (0.61)	0.35 (0.70)	0.50 (0.59)	-0.20 (0.54)	-0.06 (0.59)	0.34 (0.67)	0.18 (0.52)
Ideologically Connected Cabinet	-0.41** (0.14)	-0.29 (0.18)	-0.39*** (0.10)	-0.40** (0.15)	-0.39* (0.17)	-0.53* (0.21)	-0.43** (0.16)	-0.18 (0.22)
EU Member	-1.1* (0.54)	-0.75 (0.47)	-0.37 (0.51)	-0.39 (0.37)	-0.59 (0.48)	-0.60 (0.41)	-0.29 (0.53)	0.10 (0.48)
Refugee Benefits	-0.38 (0.92)	0.04 (0.80)	-0.77 (0.97)	0.28 (0.97)	-0.94 (1.0)	-0.05 (1.1)	0.33 (0.94)	0.36 (0.74)
Count Prev CAMs	-1.7* (0.86)	-1.7* (0.85)	-0.99** (0.33)	-1.0** (0.33)	-0.55*** (0.12)	-0.56*** (0.12)	0.006 (0.09)	-0.01 (0.09)
Conservative × Below Mean to -1 SD	-0.34 (0.63)		-1.1* (0.54)		-0.36 (0.57)		-0.09 (0.67)	
Soc Democratic × Below Mean to -1 SD	0.39 (0.65)		-0.69 (0.36)		-0.19 (0.49)		0.97* (0.46)	
Socialist × Below Mean to -1 SD	-0.76 (0.73)		-15.2*** (0.79)		-1.6** (0.55)		-0.30 (0.56)	
Mean to 1 SD		0.24 (0.59)		-0.79 (1.1)		-14.1*** (0.42)		-14.2*** (0.43)
Conservative × Mean to 1 SD		0.54 (0.72)		0.90 (1.2)		13.6*** (0.90)		15.3*** (0.67)
Soc Democratic × Mean to 1 SD		-1.0 (0.92)		0.04 (1.2)		13.4*** (0.82)		13.2*** (0.74)
Socialist × Mean to 1 SD		2.1** (0.64)		16.4*** (1.6)		16.4*** (0.67)		15.7*** (0.63)
Observations	5,292	5,292	5,341	5,341	5,341	5,341	5,291	5,291
Squared Correlation	0.07345	0.07788	0.02820	0.02983	0.01770	0.02133	0.01523	0.01786
Pseudo R ²	0.09129	0.09616	0.05982	0.06249	0.04496	0.05239	0.04113	0.05091
BIC	1,822.5	1,814.2	1,857.4	1,853.0	1,856.9	1,844.7	1,857.3	1,841.1
Country fixed effects	✓	✓	✓	✓	✓	✓	✓	✓
Decade fixed effects	✓	✓	✓	✓	✓	✓	✓	✓

Clustered (Country) standard-errors in parentheses

Signif. Codes: ***: 0.001, **: 0.01, *:0.05, .: 0.1

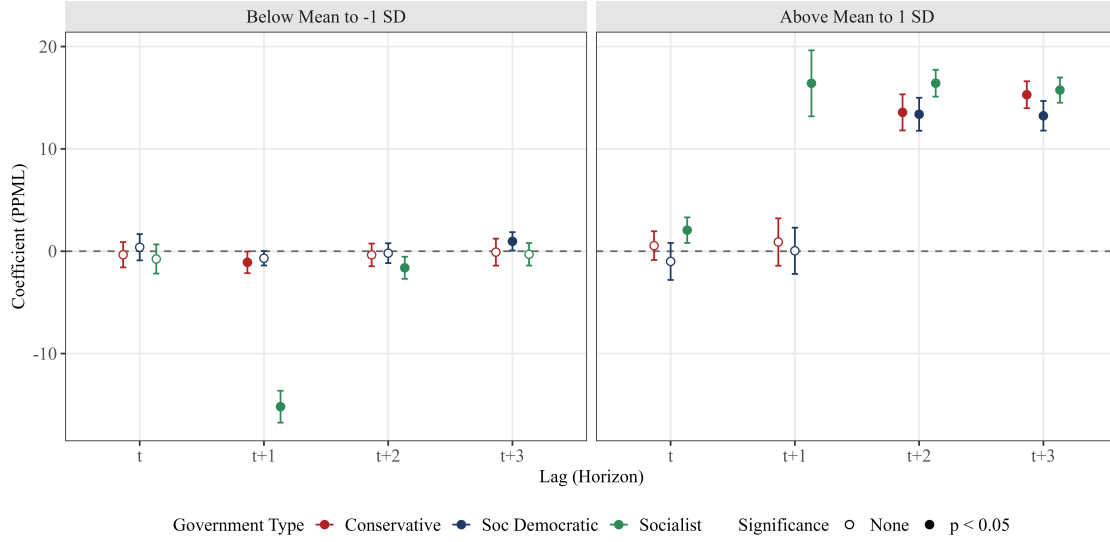


Figure B.11: Coefficient estimates for the RWP seat-share bin interactions, by government type (Conservative, Soc Democratic, Socialist), across the contemporaneous period and three annual leads. Filled points denote significance at $p < 0.05$; open points denote non-significance. Coefficients in the *Above Mean to 1 SD* panel from $t+2$ onward should be interpreted with the caution described in the text.

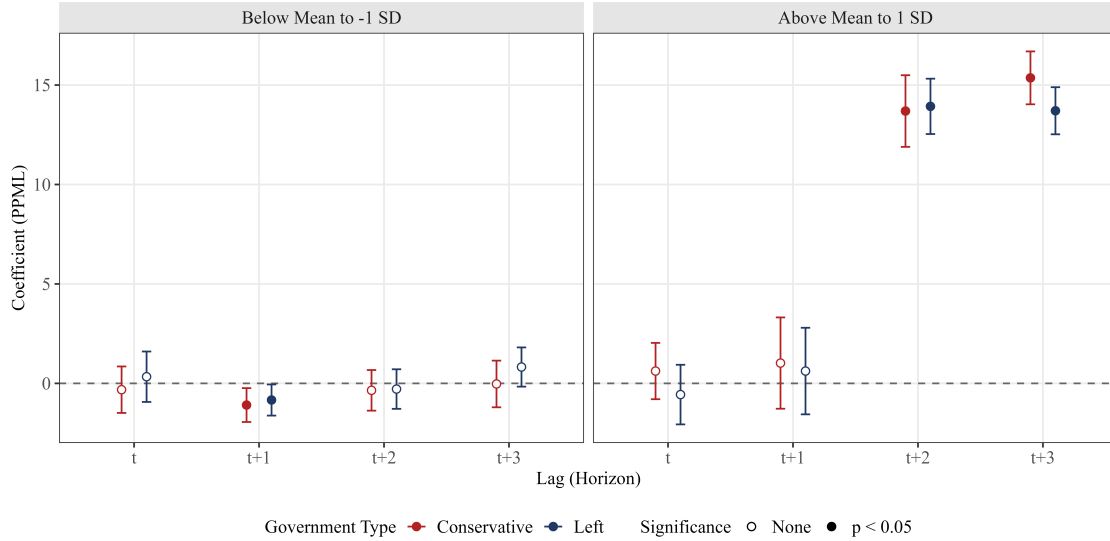


Figure B.12: Coefficient estimates as in Figure B.11, with Socialist and Soc Democratic governments collapsed into a single *Left* category. The convergence of Conservative and Left toward the same magnitude in the *Above Mean to 1 SD* bin at $t+2$ and $t+3$, despite Socialist no longer entering the estimation as a standalone category, indicates the pattern reflects the bin's sample size rather than any particular government-type coding.

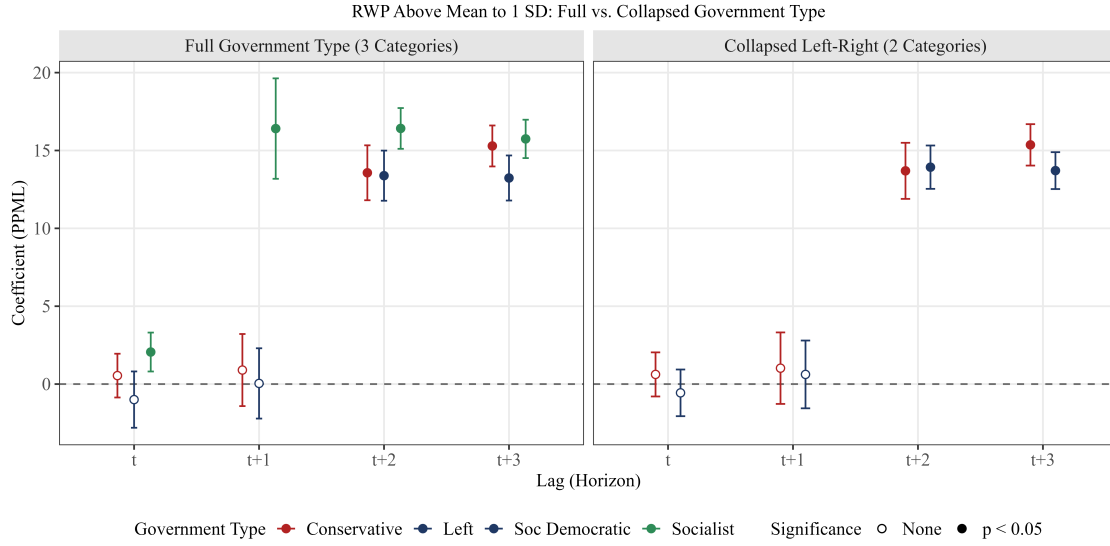


Figure B.13: *Above Mean to 1 SD* coefficients from Figures B.11 and B.12 shown side by side, comparing the full three-category government-type specification to the collapsed Left-Right specification. The same convergence toward a common magnitude at $t+2$ and $t+3$ appears in both, confirming the pattern is a property of the bin's sample size at longer horizons rather than of either government-type coding.

C Robustness Checks Using BLAs, Broader Categorization

C.1 Bilateral Labor Arrangements (BLAs) on Migration

Using data on bilateral labor arrangements (BLAs) on migration from Shrestha et al. (2025), I conduct an outcome placebo test to understand if a similar bilateral, migration related agreement happens to be similarly affected by my main independent variables in line with Eggers, Tuñón, and Dafoe (2023). For recent examples of similar placebo tests and more on their benefits refer to Peisakhin and Rozenas (2018) and Cruz and Schneider (2017). BLAs are important instruments of legal labor mobility, and Shrestha et al. (2025) show that they can increase migration by as much as 68 percent within a decade. If my main explanatory variable were simply capturing a general preference for international cooperation on migration, we would expect to see similar positive effects on BLAs.

Instead, I find estimated null and negative results across the board, where these do not match results for CAMs in main findings. In short, the interaction between government type and right-wing strength has no discernible impact on the formation of BLAs. The pattern of consistent null effects or alternating effects is exactly what a placebo test should produce. It indicates that my independent variable does not proxy for a generic cooperation, nor does it capture an overall pro-migration or anti-migration shift that would spill over into labor-mobility agreements. Rather, it affects migration security agreements specifically, strengthening confidence that the main results reflect a targeted political mechanism—right-wing pressure reshaping foreign policy on migration rather than noise or broader mobility policy trends.

Table C.9: FE PPML models measuring the count of total bilateral labor arrangements (BLAs) on migration.

	BLAs		
	(1, t)	(2, $t + 1$)	(3, $t + 2$)
Refugee Stock (Log)	0.037 (0.345)	-0.044 (0.363)	3.56*** (1.02)
Asylum Applications (Log)	0.828* (0.368)	0.809 (0.436)	1.94*** (0.558)
Conservative		-1.29 (2.16)	-68.2*** (7.13)
Soc Democratic		-0.699 (1.36)	1.83 (1.96)
Socialist		-90.1*** (3.90)	85.9 (55.7)
RWP Seat Share		0.033* (0.016)	0.418 (0.491)
Incumbent		1.00 (0.968)	7.41*** (0.801)
Conservative $\times$ RWP Seat Share		0.023 (0.052)	0.321 (0.647)
Soc Democratic $\times$ RWP Seat Share		-3.89*** (0.229)	-3.89*** (0.723)
Socialist $\times$ RWP Seat Share		4.41*** (0.237)	-5.27 (2.86)
Partner Non-Democratic			0.667*** (0.194)
Foreign Movement Index			2.06* (0.891)
Life Expectancy (in Years)			-0.063 (0.156)
Contiguous Borders			1.42 (1.52)
Distance (Log)			-0.136 (0.283)
Previous Colony			-19.9*** (2.43)
Migration Protests			-39.0*** (4.42)
GDPPC (Log)			-8.11 (16.6)
Ideologically Connected Cabinet			-29.5*** (5.16)
EU Member			64.2*** (6.77)
BLAs in Past 5 Years			4.91** (1.69)
Observations	5,629	5,629	2,624
Squared Correlation	0.00323	0.00661	0.39222
Pseudo R ²	0.07221	0.15547	0.54024
BIC	289.10	340.77	262.53
Country fixed effects	✓	✓	✓
Decade fixed effects	✓	✓	✓

Clustered (Country) standard-errors in parentheses
*Signif. Codes: ***: 0.001, **: 0.01, *: 0.05, .: 0.1*

C.2 Binary and Simple Left Measurement

Binary Measure

Table C.10 replicates the main specification using a coarse binary indicator of mainstream governance, coded 1 for Center Right, Conservative, Social Democratic, and Socialist governments and 0 for Right-Wing Populist governments, which are excluded from the main estimation sample. The binary measure is included here for completeness and transparency, but it is not the preferred operationalization for a theoretical and a methodological reason.

First, the binary variable is theoretically redundant given the PM-based categorical measure. The party family classification already expresses the mainstream vs. non-mainstream distinction with finer ideological granularity. Collapsing it into a single binary variable discards meaningful variation between left-of-center and center-right mainstream governments. Literature operationalizes government type through party family or left-right composites rather than binary mainstream indicators (Akkerman, De Lange, and Rooduijn 2016; Van Spanje 2010) as mainstream party families do not behave uniformly once disaggregated.

Second, the binary measure introduces a near-collinearity problem with country fixed effects. The binary mainstream measure assigns a value of one to all Center Right, Conservative, Social Democratic, and Socialist governments and zero otherwise. Because mainstream status is near-constant within frontier states over the sample period as most countries in the sample are governed by mainstream parties for the overwhelming majority of years the binary variable conveys, little within-country variation once country fixed effects are absorbed. For example, being France and being a mainstream government in France are almost the same thing in the data. This is manageable in the baseline and $t + 2$ specifications, where the estimation sample is stable enough for the model to separate the contributions of the binary variable and the country fixed effect. However, in the $t + 1$ specification, the lead window shifts the sample in a way that concentrates the near-collinearity over the subsample where CAMs occur, producing an implausibly large coefficient that reflects numerical instability rather than a substantive finding. Refer to the coefficient in the table. Standard diagnostics do not flag formal separation, but the jump itself is diagnostic of this underlying problem. The $t + 1$ instability is noted in the table for transparency in Table C.10 below. The PM-based categorical measures do not exhibit this instability because government type varies more meaningfully within countries over time, providing the model with genuine within-country variation to identify the interaction effect.

Table C.10: FE PPML models measuring the count of total CAMs, given simple measure of mainstream government interacted with RWP seat share.

	(1)	CAMs (2)	(3)	CAMs (t+1) (4)	CAMs (t+2) (5)
Refugee Stock (Log)	-0.02 (0.13)	-0.02 (0.15)	-0.04 (0.19)	-0.05 (0.16)	0.14 (0.21)
Asylum Applications (Log)	0.06 (0.04)	0.08 (0.04)	0.11 (0.08)	0.01 (0.10)	0.14 (0.09)
Mainstream Government		-0.20 (0.28)	-0.25 (0.25)	-0.10 (0.26)	0.03 (0.30)
RWP Seat Share		-0.01 (0.010)	-0.02 (0.01)	-0.006 (0.01)	0.0003 (0.01)
Incumbent		-0.18 (0.19)	-0.14 (0.19)	-0.13 (0.19)	-0.25 (0.19)
Mainstream Government $\times$ RWP Seat Share		0.005 (0.008)	0.01* (0.005)	0.008 (0.01)	-0.02 (0.01)
Partner Non-Democratic			-0.24 (0.14)	-0.30*** (0.07)	-0.04 (0.18)
Foreign Movement Index			-0.10* (0.05)	-0.11 (0.06)	-0.05 (0.03)
Life Expectancy (in Years)			-0.01 (0.01)	-0.02*** (0.006)	-0.007 (0.01)
Contiguous Borders			-0.24 (0.23)	-0.16 (0.26)	-0.29 (0.28)
Distance (Log)			-0.35*** (0.08)	-0.21*** (0.06)	-0.07 (0.05)
Previous Colony			0.26 (0.59)	0.01 (0.28)	-0.12 (0.12)
Migration Protests			-0.10** (0.04)	-0.17* (0.07)	-0.24*** (0.03)
GDPPC (Log)			1.1 (0.64)	0.56 (0.61)	-0.03 (0.58)
Ideologically Connected Cabinet			-0.30* (0.14)	-0.33 (0.17)	-0.27 (0.18)
EU Member			-0.96 (0.68)	-0.64 (0.52)	-0.59 (0.53)
Refugee Benefits			-0.22 (0.92)	0.009 (0.92)	-0.40 (1.0)
Count Prev CAMs			-1.3 (1.0)	-0.73 (0.52)	-0.54*** (0.11)
Observations	8,458	8,458	5,823	5,872	5,872
Squared Correlation	0.01100	0.01214	0.05442	0.02250	0.01780
Pseudo R ²	0.03577	0.03820	0.07494	0.04695	0.04797
BIC	2,754.5	2,784.1	1,979.7	2,006.6	2,011.0

Clustered (Country) standard-errors in parentheses
*Signif. Codes: ***: 0.001, **: 0.01, *: 0.05, .: 0.1*

Simple Left Measure

The collapsed left-right measure pools Socialist and Social Democratic governments into a single *Left* category while retaining Conservative and Center Right as distinct categories, with Center Right as the reference. Right-wing populist and Special governments are excluded from the estimation sample, which is consistent with the main specification. Conservative governments are preserved as a separate category rather than folded into a generic right bloc because the main results in the main tables show that the Conservative $\times$ RWP interaction, while positive, is substantively weaker than the interactions estimated for left-of-center government types. By collapsing Conservative with Center Right, I would obscure this gradient and artificially sharpen the left-right contrast. The collapsed measure therefore provides a conservative test of the second hypothesis if the left-of-center co-optation effect survives the pooling of Socialist and Social Democratic governments. Results show that despite the heterogeneity documented in the main table, it provides additional confidence that the findings hold and are not an artifact of the small Socialist cell, the PPML model, or of the specific party family coding. Table C.11 is below.

Table C.11: FE PPML models measuring the count of total CAMs, collapsed left-right measure interacted with RWP seat share.

	(1)	CAMs (2)	(3)	CAMs (t+1) (4)	CAMs (t+2) (5)
Refugee Stock (Log)	-0.06 (0.16)	-0.02 (0.14)	0.03 (0.20)	0.04 (0.17)	0.10 (0.22)
Asylum Applications (Log)	0.07 (0.05)	0.10* (0.05)	0.12 (0.07)	-0.01 (0.10)	0.12 (0.10)
Conservative		-0.56 (0.35)	-0.68 (0.61)	-0.81 (0.44)	-0.70 (0.54)
Left		0.06 (0.19)	-0.28 (0.31)	-0.40 (0.35)	-0.07 (0.46)
RWP Seat Share		-0.01 (0.008)	-0.05* (0.02)	-0.04 (0.02)	-0.07* (0.03)
Conservative $\times$ RWP Seat Share		0.02* (0.009)	0.05* (0.02)	0.06** (0.02)	0.06* (0.02)
Left $\times$ RWP Seat Share		-0.003 (0.010)	0.04* (0.02)	0.04* (0.02)	0.03 (0.02)
Partner Non-Democratic			-0.30* (0.14)	-0.30** (0.10)	-0.20* (0.10)
Foreign Movement Index			-0.08 (0.05)	-0.08 (0.06)	-0.05 (0.04)
Life Expectancy (in Years)			-0.009 (0.02)	-0.02* (0.009)	-0.01* (0.007)
Contiguous Borders			-0.17 (0.30)	-0.22 (0.33)	-0.29 (0.31)
Distance (Log)			-0.28** (0.10)	-0.23*** (0.07)	-0.05 (0.07)
Previous Colony			0.02 (0.42)	-0.10 (0.19)	-0.19 (0.11)
Migration Protests			-0.09* (0.04)	-0.15** (0.06)	-0.19*** (0.04)
GDPPC (Log)			0.98 (0.74)	0.41 (0.65)	-0.24 (0.58)
Ideologically Connected Cabinet			-0.45*** (0.11)	-0.48*** (0.12)	-0.47** (0.18)
EU Member			-0.70 (0.76)	-0.38 (0.58)	-0.08 (0.61)
Refugee Benefits			-0.67 (0.96)	-0.96 (0.96)	-1.1 (1.0)
Count Prev CAMs			-1.6 (0.85)	-1.0** (0.33)	-0.58*** (0.11)
Observations	7,745	7,745	5,292	5,341	5,341
Squared Correlation	0.01180	0.01464	0.06773	0.02876	0.01972
Pseudo R ²	0.03715	0.04263	0.09096	0.05987	0.05408
BIC	2,539.9	2,571.0	1,806.0	1,840.2	1,824.7
Country fixed effects	✓	✓	✓	✓	✓
Decade fixed effects	✓	✓	✓	✓	✓

Clustered (Country) standard-errors in parentheses

*Signif. Codes: ***: 0.01, **: 0.05, *: 0.1*

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